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Validity of the Modified Child Psychopathy Scale for Juvenile Justice Center Residents

Bruno Verschuere · Ingrid Candel · Lique Van Reenen · Andries Korebrits

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Abstract Adult psychopathy has proven to be an important clinical and forensic construct, but much less is known about juvenile psychopathy. In the present study, we examined the construct validity of the self report modified Child Psychopathy Scale mCPS; Lynam (Psychological Bulletin 120:(2), 209–234, 1997) in a sample of 57 adolescents residing in a Dutch juvenile justice center, aged between 13 and 22 years. The mCPS total score was reliably related to high externalizing problems, low empathy, high anger and aggression, high impulsivity, high (violent) delinquency, and high alcohol/drug use. Unique relations were found for the antisocial-impulsive (mCPS Factor 2), but not the callous-unemotional facet of psychopathy (mCPS Factor 1). Our findings support the validity of the mCPS in that it encompasses the antisocial-impulsive facet of psychopathy, but it is less clear whether the mCPS sufficiently captures the affective-interpersonal facet of psychopathy.

Keywords Psychopathy · Child psychopathy scale · Antisocial · Impulsivity · Aggression

Delinquency is relatively common among adolescents, but only a very small group of adolescent offenders persist in their delinquent behaviour as adults (Moffitt 1993). This small group is responsible for up to half of all crimes committed by adolescents (Farrington et al. 1986; Moffitt 1993; Wolfgang et al. 1972). Early identification of this group would create the opportunity for early intervention. Reviewing several large scale longitudinal studies, Lynam (1996) showed that boys who display antisocial behaviour and ADHD symptoms are most at risk for chronic offending. Frick and colleagues, however, cautioned against an exclusive reliance upon behavioural symptoms, and showed that the prediction of chronic offending can be further improved by including callous/unemotional traits (for a review see Frick and White 2008). This constellation of traits highly resembles a concept from the adult literature: psychopathy.

Psychopathy refers to a clinical construct encompassing both behavioural-lifestyle characteristics (e.g., impulsivity, irresponsibility, sensation seeking and chronic antisocial behaviour) and affective-interpersonal characteristics (e.g., lack of empathy, glibness, arrogance, lying, shallow affect, and lack of guilt; Cleckley 1941; Hare 2003; Patrick 2006). These two facets are encompassed by the most widely used adult psychopathy measure – the Psychopathy Checklist Revised (PCL-R; Hare 2003). PCL-R Factor1 captures the...
Psychopathy is associated with aggression and violence, and psychopathy reliably predicts general, violent, and sexual criminal recidivism (Hare et al. 2000; Quinsey et al. 2003). Irrespective of this discussion, there is a consensus that the psychopathy construct and its measures is essential before implementing the psychopathy construct and the associated risks (e.g., labelling), expected that juvenile psychopathy as indexed by the mCPS selected. Based on content and psychometric properties (i.e., item-scale correlations and internal consistency), 41 items – aimed to tap 13 of the 20 PCL items – were selected. For 7 PCL items, there were no CPS items, for example, because of the difficulties in creating an appropriate youth version of the PCL item (e.g., many short term marital relationships, revocation of conditional release, etc.). With two exceptions (i.e., lack of remorse or guilt, \( a = 0.25 \); and parasitic lifestyle, \( a = 0.36 \)), the CPS Scales had sufficient internal consistency (\( a = 0.51 - 0.76 \)), and the overall internal consistency of the CPS was very high (\( a = 0.91 \)). Data from 508 high-risk community boys from the Pittsburgh Youth study support the reliability and validity of the CPS, showing that CPS scores were associated with high self reported delinquency, high impulsivity, high externalizing problems, and low internalizing problems (Lynam 1997).

The CPS was later revised. Specifically, (1) items were added to increase the reliability and validity of several CPS Scales (e.g., the number of items in the Poverty of Affect Scale rose from 2 to 8), (2) two scales were added: Grandiosity (originally not included because of its lack of correlation with the other CPS Scales) and Boredom Susceptibility (originally not included because of difficulties in the operationalization), and (3) items tapping Criminal Versatility and Juvenile Delinquency were excluded in order to assure that the CPS is a ‘pure’ personality measure uncontaminated by antisocial behavior (for a theoretical discussion see Cooke and Michie 2001; Lilienfeld and Andrews 1996). As a result, the modified CPS (mCPS) consisted of 55 items (see e.g., Bijttebier and Decoene 2009; Spain et al. 2004). The CPS remains an instrument under development. Because of its lack of correlation with the other CPS Scales, the Grandiosity Scale was deleted again. The present study aimed to investigate the validity of the current, 50-item mCPS (see Table 1).

To examine the construct validity of the mCPS, we measured its association with theoretically relevant criterion variables: age, IQ, impulsivity, internalizing and externalizing mental problems, empathy, aggression, and delinquency. Based on the adult psychopathy literature (Blonigen et al. 2010; Hare 2003; Harpur et al. 1988; Lilienfeld and Widows 2005; Lynam 1997; Patrick 2006; Uzieblo et al. 2007), we expected that juvenile psychopathy as indexed by the mCPS total score would be associated with low age, high impulsivity, high externalizing, high disinhibition, low empathy, high aggression, and high delinquency. Furthermore, different and unique associations were expected for the two facets of psychopathy. The affective-interpersonal facet of psychopathy (i.e., mCPS Factor1) was predicted to be uniquely related to low internalizing problems, and to be more strongly associated with low empathy and high proactive aggression than the antisocial-lifestyle facet of psychopathy (i.e., mCPS Factor2). The latter factor was predicted to show unique relations with high internalizing problems, and to show a stronger relation

The self report measures of juvenile psychopathy reliably predict program failure and dropout, institutional antisocial behaviour, and general and violent criminal recidivism (Boccaccini et al. 2007; Falkenbach et al. 2003; Salekin 2008; Spain et al. 2004). Meta-analytic research on 2,867 youth aged between 14 and 19 years showed a reliable association between clinical expert ratings using the PCL: YV and general and violent criminal recidivism (Edens et al. 2007). Accumulating evidence further suggests that juvenile psychopathy, like its adult counterpart, is associated with severe and stable aggression, delinquency, and deficient processing of emotional information (Frick and White 2008). Given the controversy around the juvenile psychopathy construct and the associated risks (e.g., labelling), however, thorough validation of the juvenile psychopathy construct and its measures is essential before implementing it in forensic decision making (Salekin and Lynam 2010; Seagrave and Grisso 2002).

The present study examines the validity of the Child Psychopathy Scale. Using archival data from the Pittsburgh Youth study, Lynam tried to measure psychopathy in a similar way as the adult Psychopathy Checklist (PCL: Hare 2003) by combining items from the Child Behavior Checklist (CBCL; Achenbach 1991) and the Common-Language Q-Sort (CLQ; Caspi et al. 1992) into the Child Psychopathy Scale (CPS; Lynam 1997). Initially, a large set of 300 relevant items was
with low age, high impulsivity, high externalizing problems, high disinhibition, high reactive aggression, high alcohol/drug abuse, and high delinquency than mCPS Factor1. The present study extends previous research on the mCPS in important ways. First, we assessed the mCPS on its own (as is likely most often done in clinical and forensic settings), rather than selecting the mCPS items post hoc from the CBCL/CLQ (as was done in several CPS validation studies). This difference in assessment may affect the validity of the mCPS, for example, because the participant may more easily grasp what the mCPS aims to measure when it is assessed as a separate instrument. Such greater face validity may in turn increase the chance that the participant engages in impression management. Therefore, we found it important to examine the validity of the mCPS as a separate instrument. Second, we assessed adolescents residing in a juvenile justice center, whereas previous studies have often examined boys residing in the community. Third, as our study was completed in the Netherlands, our study extends prior work that was largely based on North-American samples.

Method

Participants

Participants were adolescents residing in the juvenile justice center “Stichting Jeugdzorg St. Joseph”, in South-Limburg of the Netherlands. Data of this study were collected for adolescents who entered the setting between December 19, 2006 and September 19, 2009. Participants completed the assessment procedure as part of the facility program. Written informed consent to use these data for research purposes was obtained from 85 participants (an estimated 75% of the population). Written parental/caregiver consent was obtained for participants below 16 years of age. mCPS data were available for 62 youth. Because we only had data for four girls, we considered only the data from the boys in our analyses (n=58). Data from one boy were excluded because of excessive missing data (see Results Section). Fifty-seven boys between 13 and 22 years of age (M=16.75, SD=1.68 years) took part in this study. Thirty-four boys (64%) were Dutch, with the remaining having a variety of different ethnic backgrounds (European [n=2], Asian [n=1], South-American [n=5], Central-American [n=1], Mid-Eastern [n=3], North-African [n=6], Black-African [n=2], Not registered [n=3]).

Material

Wechsler Intelligence Scale for Children-III/ Wechsler Adult Intelligence Scale-III (WISC-III/WAIS-III)

IQ was assessed with the Wechsler Intelligence Scale for Children - III (WISC-III; Wechsler 2003) in boys younger than 16 years of age, and with the Wechsler Adult Intelligence Scale (WAIS-III; Wechsler 1997) in those of 16 years or older.

Child Psychopathy Scale (CPS)

We used self report version of the Child Psychopathy Scale (mCPS; Lynam 1997; see http://www1.psych.purdue.edu/~dlynam/cpspage.htm). This version contains 50 yes-or-no

### Table 1 Correspondence between Lynam’s 50-Item Self Report Child Psychopathy Scale (mCPS) and Hare’s Psychopathy Checklist (PCL)

<table>
<thead>
<tr>
<th>PCL Scale</th>
<th>mCPS Scale</th>
<th>Number of items in CPS</th>
<th>mCPS Item Example</th>
</tr>
</thead>
<tbody>
<tr>
<td>Factor1</td>
<td>Glibness</td>
<td>4</td>
<td>Are you talkative?</td>
</tr>
<tr>
<td>Glibness</td>
<td>Untruthfulness</td>
<td>5</td>
<td>Will you usually tell a lie if you think you can get</td>
</tr>
<tr>
<td>Pathological Lying</td>
<td>Manipulation</td>
<td>3</td>
<td>Do you try to take advantage of other people?</td>
</tr>
<tr>
<td>Manipulative Lack of guilt</td>
<td>Lack of guilt</td>
<td>3</td>
<td>Do you usually feel guilty after doing something wrong?</td>
</tr>
<tr>
<td>Lack of remorse or guilt</td>
<td>Callousness</td>
<td>4</td>
<td>Are you considerate and thoughtful of other people?</td>
</tr>
<tr>
<td>Lack of empathy</td>
<td>Poverty of affect</td>
<td>7</td>
<td>Do you make close friendships with other people?</td>
</tr>
<tr>
<td>Shallow affect</td>
<td>Failure to accept responsibility</td>
<td>3</td>
<td>Do you try to blame other people for things that you have done?</td>
</tr>
<tr>
<td>Factor2</td>
<td>Boredom susceptibility</td>
<td>5</td>
<td>Are you easily bored?</td>
</tr>
<tr>
<td>Need for stimulation</td>
<td>Parasitic Lifestyle</td>
<td>4</td>
<td>Do you try to see how much you can get away with?</td>
</tr>
<tr>
<td>Parasitic Lifestyle</td>
<td>Lack of planning</td>
<td>3</td>
<td>Do you plan things ahead?</td>
</tr>
<tr>
<td>Lack of long term goals</td>
<td>Impulsiveness</td>
<td>3</td>
<td>Do you use your head before doing or saying something?</td>
</tr>
<tr>
<td>Impulsivity</td>
<td>Unreliability</td>
<td>3</td>
<td>Are you reliable and dependable?</td>
</tr>
<tr>
<td>Irresponsibility</td>
<td>Behavioral Dyscontrol</td>
<td>3</td>
<td>Are you easily frustrated?</td>
</tr>
</tbody>
</table>

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questions (e.g., “Are you a good liar?”). These 50 items load on 13 of the 20 PCL items (see Table 1). Moderate to high internal consistency has been reported for different versions of the CPS. Lynam (1997) found $\alpha = .91$ for the 41-item mother-reported version in a sample of about 500 boys from the Pittsburgh Youth Study. Falkenbach and colleagues (2003) found $\alpha = .77$ and $\alpha = .90$ for the 55-item self- and parent-report versions, respectively, in a sample of 69 delinquent youth from a juvenile diversion program. Bijeitbeier and Decone (2009) found internal consistency of .81, .91, and .68 of the parent-, teacher-, and self-report versions of the 55-item mCPS in a community sample of 182 children. Six-month test-retest reliability of the 58-item mCPS was .87 (Izen et al. 2010).

Using confirmatory factor analyses (CFA) on the original 41-item CPS, Lynam (1997) initially found adequate fit for the “classic” two-factor structure. This two-factor structure also produced better fit than a one factor model for the 55-item version (Bijtebieer and Decoene 2009), but fit was low (see also; Izen et al. 2010). To be consistent with Lynam (1997) and Bijtebieer and Decoene (2009), and because of extensive research on the 2-factor model in the adult psychopathy literature, we used the 2-factor model.

**The Aggression Questionnaire (AQ)**

The Aggression Questionnaire (AQ; Buss and Perry 1992) consists of 29 items that are rated on a 5-point scale from 1 (completely disagree) to 5 (completely agree), and load on physical aggression (9 items), verbal aggression (5 items), anger (7 items), and hostility (8 items). A total aggression score can be obtained by summing the scores of the subscales. Meesters and colleagues (1996) reported $\alpha$’s of .75, .50, .70, .76, and .84 for physical aggression, verbal aggression, anger, hostility, and the total score, respectively. Six-week test-retest reliability ranged from .76 (physical aggression) to .80 (total score).

**The Beck Depression Inventory (BDI)**

The Beck Depression Inventory is one of the most widely used instruments to assess depression. The original BDI (Beck et al. 1979) was used. Participants answered how they felt during the last week on 21 items that each had four response options; for example, (1) I do not feel sad; (2) I feel sad; (3) I am sad all the time and I can’t snap out of it; (4) I am so sad or unhappy that I can’t stand it. Beck and colleagues (Beck et al. 1988) reported a mean $a$ of .81 and a test-retest reliably (from 1 to 6 h to 4 months) ranging between .48 and .86.

**The Barratt Impulsiveness Scale (BIS-11)**

The Barratt Impulsiveness Scale (BIS-11; Patton et al. 1995) is widely used self-report measure of impulsivity. It consists of 30 items that are rated on a 1 (seldom or never) to 4 (nearly always or always) scale. Eight items load on Attentional Impulsivity, 11 items on Motor Impulsivity, and another 11 items on Nonplanning Impulsivity. Patton and colleagues reported $\alpha$’s ranging from .79 to .83 in four different populations.

**The Interpersonal Reactivity Index (IRI)**

The Interpersonal Reactivity Index (IRI; Davis 1983) consists of 28 items scored from 0 (completely does not fit me) to 4 (completely fits me) that tap into four 7-item subscales: Perspective Taking (PT), Empathic Concern (EC), Fantasy (F), and Personal Distress (PD). Following the author’s recommendations, no total score is calculated. The IRI measures both cognitive and affective empathy: the tendency to spontaneously adopt the view of other (PT), to experience compassion (EC) or distress in response to observed distressed (PD), and to transpose oneself into fictional situations (F). Davis reported $\alpha$’s ranging from .70 to .78, and a 60–75 day test-retest reliability ranging from .61 to .82.

**Reactive-Proactive Aggression Questionnaire (RPQ)**

The Reactive-Proactive Aggression Questionnaire (RPQ; Raine et al. 2006) consists of 23 items that are scored 0 (never), 1 (sometimes) or 2 (often). Eleven items measure aggression in response to frustration (Reactive Aggression) and 12 items measure instrumental aggression that serves to achieve one’s goals (Proactive Aggression). A total Aggression score is obtained by summing the scores of the two subscales. Raine and colleagues found $\alpha$’s of .81, .84, and .90, for Reactive Aggression, Proactive Aggression, and the total score, respectively.

**Self Reported Delinquency (SRD)**

The Self Reported Delinquency scale (SRD) was a 34-item version of Elliott’s SRD measure (Elliott and Ageton 1980) that is answered with yes-or-no responses. The SRD examines the occurrence of a variety of delinquent behaviour, including purposeful damage of goods, gang fights, weapon use, and selling and using drugs. We calculated a sum score for violent delinquency (9 items), selling and using alcohol/drugs (8 items), and a total score (34 items).

**Youth Self Report (YSR)**

The original version of the Youth Self Report (YSR; Achenbach 1991) was used. This 112-item self-report instrument assesses emotional and behavioral problems (Social Problems, Thought Problems, Attention Problems, Aggressive Behavior, Delinquent Behavior, Withdrawn,

Results

Missing Data

Missing data were handled following the procedure proposed by Lilienfeld and Widows (2005). First, data from measures with more than 20% missing data were omitted from analyses. There was a varying number of missing data, but there was data for a minimum of 44 boys for all analyses. Second, missing data from scales with less than 20% missing data were replaced by the mean of the completed items.

Descriptive Statistics

Means, standard deviations, and Chronbach’s alpha’s for all measures can be found in Table 2. The internal consistency of the mCPS factor and total scores are acceptable to good (Barker et al. 1994). mCPS total and factor scores did not differ significantly with ethnicity (Dutch vs. non-Dutch), $t’s < 1$.

mCPS Total Score

The correlations between the mCPS total and factor scores and theoretically relevant measures are reported in Table 3. The mCPS total score was negatively associated with age, IRI perspective taking, IRI empathic concern, and positively related to YSR internalizing, externalizing and total problems, BIS-11 motor, non-planning, attentive and total impulsivity, AQ physical and verbal aggression, anger and hostility, RPQ reactive and proactive aggression, and SRD alcohol/drugs, violent and total delinquency. Most correlations were moderate to large.

mCPS Factor Scores

Inspection of Table 3 shows that, with very few exceptions, bivariate correlations between mCPS factor scores and the criterion-related variables were similar to those between the mCPS total score and the criterion-related variables. The bivariate correlations did not provide evidence for differential associations of the mCPS factors scores with the criterion-related variables. This is perhaps not that surprising given that the two CPS factor scores were very highly correlated, $r = .76$, $p < .001$. Following the recommendations by Meng and colleagues (1992), we used Steiger’s Z-test to statistically compare the bivariate correlations for the two mCPS factors. mCPS Factor2 correlated more strongly with AQ Anger, $Z = 2.45$, $p < .05$, BIS-11 Attentive Impulsivity, $Z = 2.04$, $p < .05$, and Bis-11 total, $Z = 2.91$, $p < .01$, than mCPS Factor1.

<table>
<thead>
<tr>
<th>Measure</th>
<th>$M (SD)$</th>
<th>$\alpha$</th>
</tr>
</thead>
<tbody>
<tr>
<td>IQ</td>
<td>88.52 (11.64)</td>
<td>.72</td>
</tr>
<tr>
<td>mCPS Factor1</td>
<td>2.07 (1.00)</td>
<td>.72</td>
</tr>
<tr>
<td>Factor2</td>
<td>1.58 (1.09)</td>
<td>.81</td>
</tr>
<tr>
<td>Total</td>
<td>3.65 (1.96)</td>
<td>.86</td>
</tr>
<tr>
<td>YSR Internalizing</td>
<td>10.47 (6.99)</td>
<td>.84</td>
</tr>
<tr>
<td>Externalizing</td>
<td>15.90 (9.91)</td>
<td>.91</td>
</tr>
<tr>
<td>Total</td>
<td>42.35 (21.71)</td>
<td>.93</td>
</tr>
<tr>
<td>BDI</td>
<td>8.67 (6.84)</td>
<td>.78</td>
</tr>
<tr>
<td>IRI Perspective Taking</td>
<td>12.52 (5.26)</td>
<td>.71</td>
</tr>
<tr>
<td>Empathic Concern</td>
<td>16.64 (5.33)</td>
<td>.70</td>
</tr>
<tr>
<td>Fantasy</td>
<td>9.37 (5.71)</td>
<td>.73</td>
</tr>
<tr>
<td>Personal Distress</td>
<td>8.17 (5.42)</td>
<td>.72</td>
</tr>
<tr>
<td>BIS-11 Motor Impulsivity</td>
<td>21.07 (3.69)</td>
<td>.51</td>
</tr>
<tr>
<td>Non-planning Impulsivity</td>
<td>24.67 (4.88)</td>
<td>.71</td>
</tr>
<tr>
<td>Attentive Impulsivity</td>
<td>15.70 (3.52)</td>
<td>.58</td>
</tr>
<tr>
<td>Total</td>
<td>61.46 (9.50)</td>
<td>.78</td>
</tr>
<tr>
<td>AQ Physical Aggression</td>
<td>27.89 (8.75)</td>
<td>.86</td>
</tr>
<tr>
<td>Verbal Aggression</td>
<td>15.27 (2.70)</td>
<td>.40</td>
</tr>
<tr>
<td>Anger</td>
<td>17.70 (5.26)</td>
<td>.70</td>
</tr>
<tr>
<td>Hostility</td>
<td>19.79 (5.94)</td>
<td>.77</td>
</tr>
<tr>
<td>Total</td>
<td>80.66 (17.97)</td>
<td>.89</td>
</tr>
<tr>
<td>RPQ Reactive Aggression</td>
<td>10.83 (4.49)</td>
<td>.85</td>
</tr>
<tr>
<td>Proactive Aggression</td>
<td>5.90 (4.94)</td>
<td>.88</td>
</tr>
<tr>
<td>Total</td>
<td>16.73 (8.83)</td>
<td>.92</td>
</tr>
<tr>
<td>SRD Alcohol/Drugs</td>
<td>3.61 (2.12)</td>
<td>.77</td>
</tr>
<tr>
<td>Violent</td>
<td>3.86 (2.49)</td>
<td>.77</td>
</tr>
<tr>
<td>Total</td>
<td>14.02 (7.87)</td>
<td>.92</td>
</tr>
</tbody>
</table>

mCPS Child Psychopathy Scale; YSR Youth Self Report; BDI Beck Depression Inventory; IRI Interpersonal Reactivity Index; BIS-11 Barratt Impulsiveness Scale 11; AQ Aggression Questionnaire; RPQ Reactive-Proactive Aggression Questionnaire; SRD Self Reported Delinquency

Table 2 Mean (M; with Standard deviation, SD) and internal consistency (Chronbach’s $\alpha$) of all measures included in the correlation analyses
Because suppressor effects have been reported between psychopathy factor scores in the prediction of criterion-related variables (Blonigen et al. 2010), we also calculated partial correlations. Partial correlations shed light on the unique relationship between the factor scores and the criterion-related variables. After controlling for mCPS Factor 2, none out of 18 (0%) significant bivariate correlations for mCPS Factor 1 remained significant. After controlling for mCPS Factor 1, 14 out of 20 (70%) significant bivariate correlations for mCPS Factor 2 remained significant. This suggests that mCPS Factor 2 rather than Factor 1 uniquely relates to the criterion-related variables. Analyses with the rationally derived, but untested, 3-factor structure (Dividing Factor 1 into an Affect and an Interpersonal factor, and excluding Behavioral Dyscontrol from Factor 2; Spain et al. 2004) showed a similar, yet less dramatic, pattern. After controlling for the other two factors, the number of significant correlations dropped from 17 to 5 for mCPS Affect (correlations with YSR externalizing, BIS-11 Non-planning Impulsivity, AQ Physical Aggression, AQ total, and RPQ total remained significant), and from 17 to 2 for mCPS Interpersonal (correlations with YSR internalizing and YSR total remained significant; in addition IRI Fantasy became significant). For mCPS Lifestyle, however, 11 out of 20 correlations remained significant after controlling for the other two factors (correlations with YSR externalizing,
YSR total, BIS-11 Motor Impulsivity, BIS-11 Non-planning Impulsivity, BIS-11 total, AQ anger, AQ total, RPQ reactive aggression, RPQ proactive aggression, and RPQ total remained significant).

Discussion

Researchers have expressed great concern with regard to the concept of juvenile psychopathy (Hart et al. 2002). Several characteristics of psychopathy may be too common among adolescents and reflect transient developmental states rather than stable personality traits. Moreover, labeling of youth as psychopathic, accurate or not, can have great consequences. On the other hand, accurate identification of juvenile psychopathy could have great benefits, both for the individual and for society. Thus, thorough validation of juvenile psychopathy measures is required before applying them in juvenile forensic settings (Hart et al. 2002; Seagrave and Grisso 2002). Seagrave and Grisso argued that it needs to be firmly established that juvenile psychopathy (1) can be measured reliably, (2) is similar to its adult counterpart, and (3) is stable across time. The present study tackles the first two aspects, by assessing the modified Child Psychopathy Scale along with a range of criterion-related variables in a sample of 57 juvenile justice center residents.

We found that the internal consistency of the mCPS was moderate for mCPS Factor1 (α = .72), and good for mCPS Factor2 (α = .81) and mCPS Total score (α = .86). In a community sample of 182 Flemish adolescents, Bijttebier and Decoene (2009), however, found that the internal consistency of the self-report mCPS (Factor1: α = .54, Factor2: α = .50, Total: α = .68) was not sufficient (i.e., lower than .70), and lower than that scored by the parent (Factor1: α = .70, Factor2: α = .70, Total: α = .81) or the teacher (Factor1: α = .81, Factor2: α = .85, Total: α = .91). Age differences may account for these findings. Whereas our sample had a mean age of 16.75 years, the Flemish community sample had a mean age of 13.20 years and younger children may have had difficulty understanding the questions. Indeed, the same group (Roose et al. 2010) found somewhat better internal consistency in 455 Flemish adolescents with a mean age of 16.67 years (Factor1: α = .67, Factor2: α = .64, Total: α = .77). Furthermore, the internal consistency of the current 50-item version of the mCPS may be better than the 55-item version used in the Flemish community samples. Finally, our sample (juvenile boys) may better resemble the sample in which the mCPS was developed: boys at high risk for delinquency. Douglas and colleagues (Boccaccini et al. 2007; Douglas et al. 2008; Spain et al. 2004) found moderate to good internal consistency for mCPS total score (α = .78 to .87) in different samples of adolescent offenders with a mean age around 15–16 years. Taken together, the data indicate that the internal consistency of the self report mCPS is adequate in male adolescent forensic samples, but not necessarily in other samples (e.g., girls, younger children, the community).

As predicted, the mCPS total score was associated with low age, low empathy (IRI), high externalizing (YSR), high impulsivity (BIS-11), high aggression (AQ, RPQ), and high delinquency (SRD). The associations were moderate to large, with most correlations in the .30 to .70 range. Of particular interest for forensic decision making is the association between juvenile psychopathy on the one hand and aggression and (violent) delinquency on the other hand. As such, juvenile psychopathy, as assessed with the mCPS, indeed seems to look like its adult counterpart (Hare 2003; Harpur et al. 1988; Lilienfeld and Andrews 1996; Uzieblo et al. 2007). Two findings qualify this finding.

First, the mCPS was positively related to YSR internalizing problems. This is unexpected, as it would indicate that the higher the psychopathy score, the more neurotic, anxious, worried, and stressed the adolescent is (see also Lee et al. 2010). This unexpected association is not an isolated finding. Lynam (1997) observed a positive correlation between the mCPS and internalizing problems. Similarly, Salekin and colleagues (2005) observed a positive association between neuroticism and juvenile psychopathy as assessed by both the mCPS and the APSD. Salekin and colleagues reasoned that the behavioral consequences of juvenile psychopathy (e.g., incarceration) may cause anxiety, and they regarded anxiety as a positive sign that may provide a useful starting point for treatment. Although we would like to embrace this intriguing and optimistic possibility, it needs to be admitted that it is post hoc and not testable due to the correlational nature of the data.

Second, whereas the data for the mCPS total score support the validity of the juvenile psychopathy concept, the results for the mCPS factor scores are not so much in line with the adult psychopathy literature. We did not find the predicted differential and unique pattern of relations for the two mCPS factors. Based upon the adult literature, we expected that mCPS Factor1 would be inversely related to internalizing and negligibly related to externalizing. We also expected that mCPS Factor2 would be positively related to both externalizing and neuroticism. Predictions were confirmed for mCPS Factor2, but not for mCPS Factor1. It has previously been argued that suppressor effects might obscure these effects (Benning et al. 2005; Blonigen et al. 2010). However, accounting for suppressor effects using partial correlations did not reveal the expected effects for mCPS Factor1. Taken together, the mCPS encompasses the antisocial-impulsive facet of psychopathy, but it is less clear whether it sufficiently captures the affective-interpersonal facet of psychopathy. A more extensive and a more indirect (i.e., Benning et al. 2005; Birbaumer et al. 2005; Frick and
White (2008; Patrick 1994) assessment of affective-interpersonal psychopathy features may clarify this issue.

This study is not without its limitations. First, with a minimum of 44 participants in the analyses, sample size was modest. Second, and most notably, the present findings are based upon self-report measures. Although these measures were well-validated and may provide unique information (Lilienfeld and Andrews 1996), a downside is that psychopathy is associated with lack of insight (Hare 2003) and is subject to presentation biases and strategies. To that extent it is unfortunate that we had no other-report or outcome measures (e.g., rule-breaking during treatment or recidivism). Despite these limitations, our study indicates that (1) the internal consistency of the self report mCPS is adequate in male adolescent forensic samples, and that (2) the mCPS total score was reliably related to the high externalizing problems, low empathy, high anger and aggression, high impulsivity, high (violent) delinquency, and high alcohol/drug use, yet (3) question whether the mCPS sufficiently captures the affective-interpersonal facet of psychopathy.

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References


