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How Much Time Do You Spend Online? Understanding and Improving the Accuracy of Self-Reported Measures of Internet Use

Theo Araujo, Anke Wonneberger, Peter Neijens, and Claes de Vreese

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ABSTRACT

Given the importance of survey measures of online media use for communication research, it is crucial to assess and improve their quality, in particular because the increasingly fragmented and ubiquitous usage of internet complicates the accuracy of self-reported measures. This study contributes to the discussion regarding the accuracy of self-reported internet use by presenting relevant factors potentially affecting biases of self-reports and testing survey design strategies to improve accuracy. Combining automatic tracking data and survey data from the same participants (N = 690) confirmed low levels of accuracy and tendencies of over-reporting. The analysis revealed biases due to a range of factors associated with the intensity of (actual) internet usage, propensity to multitask, day of reference, and the usage of mobile devices. An anchoring technique could not be proved to reduce inaccuracies of reporting behavior. Several recommendations for research practice follow from these findings.

With a lot of media use shifting from traditional to online media environments, measures of online media behavior receive a prominent role in communication research. Users of digital media leave an abundance of traces which could, at least in theory, provide accurate measures of online media use without relying on self-reports. Access to these highly accurate tracking data, however, is limited due to concerns regarding data privacy, technical challenges and high costs of retrieving, storing, and processing these data. This is why self-reported measures of online media use remain crucial for many research areas including political (Kruikemeier, van Noort, Vliegenthart, & de Vreese, 2014), entertainment (Kardefelt-Winther, 2014), marketing (Jiang, Yang, & Jun, 2013), or health communication (Xiao, Sharman, Rao, & Upadhyaya, 2014), as well as studies about media use in general (Voorveld, Segijn, Ketelaar, & Smit, 2014).

Online behavior—increasingly fragmented and scattered across situations, devices, and platforms—poses a critical challenge for the accuracy of self-reports (de Vreese & Neijens, 2016). Studies comparing self-reported measures with tracking data—that can be considered a “gold standard” for the measurement of online behavior—indeed indicate that respondents are, in general, not accurate when providing estimates for their online media use. This is true for frequency of internet use in general (Scharkow, 2016), social network sites (Junco, 2013; Scharkow, 2016), and websites (Revilla, Ochoa, & Loewe, 2016). Because self-reported measures of online media use are expected to remain a crucial instrument for communication research, it is important to go beyond knowing the level of (in)accuracy of these measures. An important first step is to understand the factors that influence under- and over-reporting of online media use. Understanding these factors allows researchers to interpret self-reports of online media use more accurately, and to correct the respondent’s estimates.
for known biases brought in by such factors. The first aim of the present study is, therefore, to present a systematic account of factors influencing biases of self-reports.

A second, and perhaps even more critical, step is to establish which survey design strategies can be adopted to aid respondents in their estimations, and therefore improve the accuracy of such measures.

The present study combines self-reported measures of internet use with tracking data in order to address these challenges, and contributes to communication research in several ways. First, the present study offers a systematic approach to factors potentially affecting the accuracy of self-reports of online behavior by including personal characteristics and media-use related variables. In doing so, this study extends earlier findings regarding the impact of sociodemographics and frequency of internet use on under- and over-reporting of online media use (Scharkow, 2016) by uncovering the effects of factors yet unexplored such as multitasking, interest in the survey, and tablet use. Second, the present study employs an experimental design to test how different survey strategies improve the accuracy of self-reported measures of online media use. Finally, this study provides a set of actionable guidelines for communication researchers to employ in studies that require self-reported measures of online media use.

**Factors influencing response accuracy**

Systematic differences in response behavior related to specific characteristics of respondents present a severe threat to the accuracy of survey measures (Prior, 2009b; Southwell et al., 2010). Previous research has mainly looked at biases of self-reports due to sociodemographic factors (Prior, 2009b; Scharkow, 2016). However, these factors alone do not offer explanations for differences in the accuracy of self-reports. These differences may actually be caused by various motivational or cognitive processes that are related to individual characteristics but also to factors related to media use. Specifically, factors related to the fragmented and ubiquitous nature of internet use may potentially affect the quality of reporting behavior.

The respondent’s involvement with the survey as a whole could as an individual factor be responsible for systematic differences in the accuracy of self-reports. As a motivational factor, a respondent’s interest in the survey may be related to satisficing, that is, tendencies to reduce the cognitive effort that is required to retrieve the necessary information or make a valid estimation of one’s behavior (Krosnick, 1991). The level of interest in the survey may influence the level of accuracy in self-reports, with the underlying assumption that respondents who are more motivated will potentially make a higher effort to accurately recall their behavior and consequently show less tendencies of satisficing.

H1: The higher the level of interest in the survey, the higher the accuracy of self-reports.

Individual media-use behavior might affect the accuracy of self-reports. Respondents who use the internet more frequently, or who have a stronger tendency to multitask, may have more difficulty in recalling their specific behavior. Low-frequency events are generally more easily remembered correctly than more frequent events (Schwarz & Oyserman, 2001).

In addition, research has shown that media multitasking—performing several media- and non-media related tasks simultaneously—decreases the attention that is devoted to each activity (Duff, Yoon, Wang, & Anghelcev, 2014; Voorveld, 2011) and is associated more generally with lower performance levels (Rubinstein, Meyer, & Evans, 2001; Segijn, Voorveld, Vandeborg, & Smit, 2017). Consequently, recalling web use may be less accurate if the internet is used more often in multi-tasking situations.

H2: The higher the level of internet use, the lower the accuracy of self-reports.

H3: The higher the level of multitasking, the lower the accuracy of self-reports.
While the likelihood for media-multitasking can be explained by sociodemographic factors, such as age and gender, as well as personality traits (Duff et al., 2014), also the context of Internet use has been found important. Specifically, mobile media devices encourage rapid and frequent task-switching (Rubinstein et al., 2001; Zhong, 2013). So did 90% of tablet users report to engage in other activities while using their tablets (Moses, 2012). Accordingly, the use of tablets or other mobile media devices may also decrease the accuracy of self-reports.

H4: The accuracy of self-reports is lower for the use of mobile devices compared to PCs.

Improving the accuracy of self-reports through survey design

Improving survey design strategies is an important way to improve accuracy of self-reports. One possibility is to include specific reference periods in survey questions, or to ask respondents to estimate their typical behavior (Chang & Krosnick, 2003; Price, 1993; Wonneberger & Irazoqui, 2016). Earlier research highlights tradeoffs when it comes to using either of these options, i.e., asking about a specific (recent) time period (e.g., last week), or asking about typical behavior (e.g., an average week). On the one hand, asking about a specific time period in the recent past potentially reduces the cognitive load on respondents, and brings more accurate responses. This recency effect has been found, for example, for estimation of TV exposure (Wonneberger & Irazoqui, 2016). On the other hand, earlier research also found that respondents tend to over-report typical behavior and may actually under-report their behavior when being asked about more recent time periods indicating that asking about a typical week may have higher predictive validity for other outcomes, such as current events knowledge (Althaus & Tewksbury, 2007; Chang & Krosnick, 2003). An alternative explanation could be that the typical week question is confounded with attitudes, such as political interest or involvement, and therefore is more strongly correlated with knowledge (Prior, 2009a). Consequently, “typical week” would not yield more valid measures compared to more recent time periods. Considering that these two strategies are yet to be compared for internet use, we propose the following research question:

RQ1: Which reference period (recent vs. typical) is more accurate for questions regarding the duration of internet use?

Another possibility to aid respondents is to provide an anchor that facilitates their estimation (Belli, 1998; Burton & Blair, 1991; Schwarz & Oyserman, 2001). If available, population averages can provide useful cues for respondents to accurately estimate the frequency of their own behavior (Burton & Blair, 1991; Prior, 2009a; Schwarz & Oyserman, 2001). Alternatively, cues on individual situations of media use might improve response accuracy (Jerit et al., 2016; Potts & Seger, 2013). The internet is typically used at scattered moments throughout the day and oftentimes for incidental aims, such as checking e-mails or the news. This type of non-habitual or irregular behavior might complicate an accurate recall or estimation of online activities as opposed to more regular behavior that is easier to recall (Schwarz, 1999). Providing contextual cues by asking questions about one’s past has been found effective to activate memories of past behavior frequencies and to aid correct estimation (Menon & Yorkston, 2000; Potts & Seger, 2013). Such mechanism of memory activation might also facilitate more accurate recall and estimation of more recent behavior.

H5: Anchoring in a survey question improves accuracy in respondents’ estimates of internet use compared to no anchoring.
Discerning under- and overreporting

While the discussion about possible problems of self-reported measures often concerns accuracy in general, previous research revealed theoretical and empirical arguments why underreporting and overreporting should be distinguished. From a theoretical point of view, for specific types of media exposure a prevailing direction of inaccurate reporting behavior has been attributed to the perceived social desirability of the behavior (Holbrook, Green, & Krosnick, 2003; Kahn, Ratan, & Williams, 2014; Smit & Neijens, 2011). The underlying expectation is that respondents would over-report behavior that is perceived to be positive by their social group, and under-report behavior perceived to be negative. The influence of social desirability can be assessed as an individual trait (i.e., the respondent’s tendency to answer questions in a socially desirable manner) or as characteristics from different social groups (i.e., higher social classes may be expected to behave in different ways than lower social classes). Earlier research has found, for example, that members of higher socioeconomic classes are more likely to over-report their exposure to television news (Prior, 2009b) but also more likely to under-report the time they spend video gaming (Kahn et al., 2014). For general Internet use that is assessed in the present study, there are no clear expectations regarding the influence of social desirability.

However, also previous empirical findings indicate different mechanisms behind under- and overreporting. Revilla et al. (2016), for instance, found that visits of some websites were systematically under-reported while others were over-reported. The study on accuracy of self-reported internet use by Scharkow (2016) revealed that different predictors were associated to under- and overreporting indicating different underlying mechanisms for both types of errors. The following research question addresses these possible differences:

RQ2: To what extend are there differences in how personal characteristics and media-use factors relate to under- and overreporting?

Specifically, concerning the influence of the amount of media use on response accuracy previous research has yielded consistent findings for TV exposure (Wonneberger & Irazoqui, 2017), mobile phone use (Vanden Abeele, Beullens, & Roe, 2013) as well as for internet use (Scharkow, 2016). The general pattern here is that light users have a tendency to over-report while heavy users have a tendency to under-report their use of the medium in question.

H6: Lower levels of internet use are associated to higher levels of over-reporting (a) and higher levels of internet use are associated to higher levels of under-reporting (b).

Methods

Sample

This study combines automatic tracking data of online behavior and a survey on online media use of the same respondents. The sample was drawn from an online panel with approximately 200,000 respondents (50,000 households) who regularly take part in surveys of one of the largest research companies in the Netherlands, TNS NIPO. The Netherlands is an ideal case for this study, given the extremely high levels of internet penetration (above 95%).

Approximately 9,000 of the households were part of an ongoing panel that had installed software that enables tracking of internet use on PCs (laptops and desktops), and Android tablets, enabling a comparison between these devices. To take part in this specific panel, households were randomly selected from the general panel and were asked to install the software, receiving an incentive to do so. The installation was a voluntary decision, and panel participants had to give an opt-in after being informed about what type of data would be tracked. As indicated by the panel research company, the complete process followed ESOMAR rules, and the panel was already collecting tracking data prior to this study.
Out of the households that had the tracking software installed, a sample was selected that met the following conditions: (1) the tablet and/or PC must have used the internet in the previous month, (2) for PCs, the desktop or laptop should be registered for one person in the household (so that individual internet activity could be registered), and (3) the household must have been composed of only one person (to ensure that the device was only used by one person). From all 9,000 panel participants with the tracking software installed, 921 met the criteria above, and were invited to participate in the survey, yielding 690 responses (AAPOR response rate $1 = 75\%$).

The majority of the respondents (55.9\%) belonged to a sample of PC users, and the remaining belonged to a sample of tablet users. The average age was 53.99 ($SD = 14.28$), and 58\% of the respondents were female. Education was measured according to the Dutch education system using seven categories, with the lowest level being basic education (2\% of the sample), and the highest level being master- or doctorate-level (9\% of the sample).

**Procedure**

The respondents answered questions about their general internet use on the specific device (PC or tablet) in which the tracking software was installed. A split-sample experiment was integrated in the survey. Before answering the main questions, a random selection of 50\% of the respondents first answered three anchoring questions. The first two questions asked the respondent to think about the previous day/the previous week, and indicate what type of day/week it was (e.g. normal day/week, holidays, being sick, etc.). In addition, respondents were asked to think back in which situations they made use of their private computer or tablet (e.g., at home, at work, while commuting, etc.). This anchoring procedure aims at improving recall, and has been used successfully applied in earlier studies (Bronner & Neijens, 2006). The duration of Internet use was asked with respondents answering questions about their internet use on the day before (yesterday), and on an average day. For all Internet use self-reported measures, respondents were requested to provide the information specifically for the device (tablet or laptop) on which the tracking software was installed (i.e., not to report their internet use on other devices).

**Self-reported measures**

**Internet use duration**

Respondents informed how long they used the internet in number of minutes. On average, respondents reported using the internet 126.74 min ($SD = 120.45$) on the day before, and 143.14 min ($SD = 125.49$) on an average day.

**Multitasking**

Respondents answered an adapted version of a multitasking scale (Wang et al., 2012), with the items being “I multitask often when I use a PC/tablet” and “I can get more things done when I multitask”. The answers were given on a 7-point scale ($M = 3.89, SD = 1.55, \alpha = 0.80$). Respondents also answered two questions from the behavioral media multitasking scale (Duff et al., 2014), which included “How often do you multitask in general? (e.g., talk to a friend while watching TV)” and “How often do you use multiple media at the same time? (e.g., use computer while watching TV)”. The answers were also given on a 7-point scale ($M = 3.76, SD = 1.49, \alpha = 0.67$). Because both scales were highly correlated $r(688) = .57, p < .001$, only the first scale (by Wang et al., 2012) was used in the analysis given its higher reliability.

**Survey interest**

Respondents answered a single-item measure on a 7-point scale on how interested they were in the survey ($M = 5.28, SD = 1.22$).
Sociodemographic characteristics were included as control variables, considering that previous research has indicated systematic differences in response behavior due to age, gender, and education (Southwell et al., 2010). The self-reported measures are summarized in Table 1A.

**Tracking data**

The research company provided all the respondents’ tracking data from the previous month. The tracking software captured each URL accessed by the respondent, with timestamps for when the respondent first visited the URL, and the number of seconds in which the URL remained active in the browser of the respondent. For this specific tracking software, a URL is considered to be active when it is the one being displayed in the browser, meaning that other URLs that may be open in other tabs are not considered to be active. The number of active seconds is measured as the time between the URL first becoming active in the browser (i.e., displayed to the respondent) and a different URL becoming active in the browser. The same software was used for desktops and for mobile devices (tablets).

The data contained about 3 million rows, with each row containing a timestamp, the number of seconds of activity. We used a series of custom Python scripts to generate a set of variables that reflected the online behavior of each respondent. These variables always related the internet use on the device where the respondent had the tracking software installed, and were created as follows.

**Minutes of internet use (yesterday)**

The number of minutes that a respondent used the internet on the day before she or he answered the survey (aggregated for all URLs or domains). On average, respondents used the internet for 103.55 min (SD = 140.69) on the day before they answered the survey.

**Minutes of internet use (average day)**

The number of minutes for all URLs or domains used by each respondent from the moment that the tracking data started (about one month before the survey) and the day before she or he answered the survey were aggregated. The sum of active seconds was then divided by the number of days that the

<table>
<thead>
<tr>
<th>Measure</th>
<th>Definition</th>
<th>M</th>
<th>SD</th>
</tr>
</thead>
<tbody>
<tr>
<td>Survey Interest</td>
<td>Single question on how interested the respondent was in the survey (7-point scale)</td>
<td>5.28</td>
<td>1.22</td>
</tr>
<tr>
<td>Internet Use Duration</td>
<td>Number of minutes for:</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>The previous day</td>
<td>126.74</td>
<td>120.45</td>
</tr>
<tr>
<td></td>
<td>An average day</td>
<td>143.14</td>
<td>125.49</td>
</tr>
<tr>
<td>Multitasking</td>
<td>Preference for multitasking scale (Wang et al., 2012). Answers given on a 7-point scale (α = 0.80).</td>
<td>3.89</td>
<td>1.55</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Sample characteristics</th>
<th>Definition</th>
<th>M</th>
<th>SD</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age</td>
<td>Age in years</td>
<td>53.99</td>
<td>14.28</td>
</tr>
<tr>
<td>Gender</td>
<td>Male (0), Female (1)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Education</td>
<td>Measured according to the Dutch education system using seven categories</td>
<td>4.65</td>
<td>1.59</td>
</tr>
</tbody>
</table>

1Due to technological limitations, tablet measures contained only daily information by respondent aggregated at the domain level (e.g., google.com) instead of detailed URL information (e.g., google.com/mail). This domain-level information considered the start of the measurement of active seconds as the timestamp at which the domain was first active, and the end as the timestamp at which the domain was last active within a navigation session (defined as ongoing activity by a respondent, with no more than 120 sec between URLs). If a domain was used in more than one session throughout the day, all the active seconds of each session were aggregated. For some respondents, however, there was no end timestamp registered—and this was recorded as 0 active seconds. For these instances, we used the average seconds per session for that respondent on other days. Because of this limitation, all analyses related to duration are performed twice: (1) with the averages replacing the instances in which a domain has 0 active seconds, and (2) with the 0 active seconds as an actual measure. Only the first analysis is reported unless there are strong differences in the nature of the relationships between the independent and dependent variables of the study.
respondent used the internet during that period, generating a measure of the average number of minutes of internet usage per day, considering only the days on which there was internet usage. On average, respondents used the internet for 126.27 minutes ($SD = 115.93$) per day during the tracking period.

The measures related to the tracking data are summarized in Table 1B.

### Comparing self-reports with tracking data

After calculating duration and frequency of internet use across different reference periods using the tracking data, we then created three types of variables to explore the accuracy of self-reports, as outlined below.

#### Self-report absolute errors

For this type of variable, we subtracted the information coming from the tracking data for each respondent for a given measure (e.g., minutes yesterday) from its self-reported version (i.e., error = self-report — tracking data). Because we are interested for this variable in the magnitude of error, we used the absolute value of the error. The average absolute error of the duration of Internet use was 100.64 min ($SD = 121.34$) for the day before, and 94.66 min ($SD = 109.74$) for an average day.

#### Under-reporting

We considered that a respondent under-reported the behavior when the tracking data showed more activity than what was self-reported. The under-reporting variable was therefore calculated by subtracting the tracking data value from the self-reported. Instances in which the respondent had reported correctly (i.e., no error) or over-reported their behavior were set as 0 for that variable. Respondents under-reported their internet use on average by 38.72 min ($SD = 100.98$) for the day before, and by 38.89 min ($SD = 83.37$) for an average day.

#### Over-reporting

We considered that a respondent over-reported the behavior when the tracking data showed less activity than what was self-reported. The over-reporting variable was therefore calculated by subtracting the self-reported value from the tracking data value. Instances in which the respondents had reported correctly (i.e., no error) or under-reported their behavior were set as 0 for that variable. Respondents over-reported their internet use on average by 61.91 min ($SD = 96.59$) for the day before, and by 55.77 min ($SD = 97.14$) for an average day.

The measures comparing the self-reports with the tracking data are summarized in Table 1C.

### Results

#### Understanding the accuracy of self-reported measures

The first objective of this study was to understand the accuracy of self-reported measures, and the factors that influence this accuracy. As indicated in Table 1C, the average absolute error for self-reported measures of internet use duration was 100.64 min ($SD = 121.34$) for the previous day, and 94.66 min ($SD = 109.74$) for an average day. The correlations between self-reported measures and
tracking data were about 0.3 for both the yesterday and average-day measure, indicating low to moderate levels of convergent validity (see Table 2).

OLS regression models on the different types of errors were estimated to examine to what extent inaccuracies of self-reports were systematically related to individual characteristics. We consider the models on the absolute errors as conclusive regarding hypotheses 1–5 and the first research question. The more specific models on over- and under-reporting offer additional information about the prevailing nature of inaccuracies and serve to answer the second research question and test hypothesis 6. Furthermore, plots of predicted values facilitated the interpretation of significant effects found for the more specific models of over- and under-reporting.

Interest in the survey was not associated with lower levels of absolute errors for self-reported duration of internet use, not providing support for H1. Also, no associations were found for levels of over-reporting. However, respondents with higher levels of interest in the survey were less likely to under-report the minutes spent online on the previous day while no association was found for under-reporting of internet use on an average day (see Table 3 and Figures 1A–1C). Overall, we found only a very limited influence of survey interest on response accuracy.

Actual levels of internet use (as observed in the tracking data) confirmed our expectations when it comes to accuracy of self-reports. Respondents with higher levels of actual internet use duration were associated with higher levels of absolute errors for self-reports (see Table 3 and Figure 2A), providing support for H2. The more specific models of over- and under-reporting revealed tendencies corresponding to previous research. Respondents with lower levels of internet use were associated to higher levels of over-reporting, supporting H6a. In contrast, those with higher actual levels of media exposure were less likely to over-report their behavior for the previous day and an average day. In addition, those with higher actual usage levels were more likely to under-report their behavior, supporting H6b (see Table 3, Figures 2B and 2C). In sum, heavy users were more likely to under-report while light users were more likely to over-report their internet use.

<table>
<thead>
<tr>
<th>Measure</th>
<th>Definition</th>
<th>Minutes spent online (Duration)</th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Yesterday</td>
<td>Average Day</td>
<td></td>
</tr>
<tr>
<td>Self-report absolute errors</td>
<td>The information coming from the tracking data for each respondent for a given measurement (e.g., minutes yesterday) was subtracted from its self-reported version. Because we are interested in the magnitude of error, we used the absolute value of the error.</td>
<td>100.64 (121.34)</td>
<td>94.66 (109.74)</td>
<td></td>
</tr>
<tr>
<td>Under-reporting</td>
<td>The under-reporting variable was calculated by subtracting the tracking data value from the self-reported. Instances in which the respondent had reported correctly (i.e., no error) or over-reported their behavior were set as 0 for that variable.</td>
<td>38.72 (100.98)</td>
<td>38.89 (83.37)</td>
<td></td>
</tr>
<tr>
<td>Over-reporting</td>
<td>The over-reporting variable was calculated by subtracting the self-reported value from the tracking data value. Instances in which the respondent had reported correctly (i.e., no error) or under-reported their behavior were set as 0 for that variable.</td>
<td>61.91 (96.59)</td>
<td>55.77 (97.14)</td>
<td></td>
</tr>
</tbody>
</table>

Note: Standard deviations reported in parentheses.
Multitasking was not associated with absolute errors for self-reports of internet use (see Table 3 and Figure 3A). This does not provide support for H3. However, respondents with higher levels of multitasking were less likely to under-report their internet use (see Table 3, Figures 3B and 3C).

Finally, tablet users did not differ in their accuracy in terms of the levels of the absolute errors of internet use, not providing support for H4. However, the more specific models revealed that tablet users were more likely to under-report their internet use across all reference periods, and less likely to over-report their internet use (see Table 3 and Figures 4A–4C).

Concerning our second research question regarding the differences between predictors of under- and over-reporting we can conclude that multitasking was the only factor systematically related to lower levels of under-reporting internet use. Higher levels of under-reporting in combination with lower levels of over-reporting, in contrast, were related to high levels of actual internet use as well as tablet use.
Improving the accuracy of self-reports through survey design

The second objective of this study was to explore the effectiveness of different survey design strategies on improving the accuracy of self-reports. As seen in Table 3 and Figure 5A, the anchoring question is not related to absolute levels of error of self-reported internet use, not providing support for H5. In addition, under- or over-reporting were not significantly related to the use of anchor questions. Moreover, as opposed to our expectations, the predicted values displayed in Figures 5B and 5C revealed higher levels of over-reporting and lower levels of under-reporting for the anchor condition indicating that the specific anchor applied in this study did not improve accuracy.
Second, exposing respondents to different reference periods, in response to RQ2, is not associated with differences in absolute errors ($F(1,689) = 2.116$, $p = .146$), however, when removing outliers$^2$ for both absolute errors for an average day and for the previous day, the differences become significant ($F(1,667) = 4.167$, $p < .05$), with Bonferroni post-hoc adjustments indicating that respondents had higher levels of absolute errors for the previous day in comparison to an average day ($M_{\text{difference}} = 6.210$, $SE = 3.042$). The repeated-measures ANOVA for over-reporting showed significant differences ($F(1,689) = 4.149$, $p < .05$), with Bonferroni post-hoc adjustments indicating that respondents had higher levels of over-reporting for the previous day in comparison to an average day.

$^2$Cases with z-scores above 3.29 were considered outliers for this analysis.
day \((M_{\text{difference}} = 6.145, SE = 3.017)\). For under-reporting, however, the differences were not significant \(F(1,689) = .003, p = .958\).\(^3\)

Table 4 presents the key findings of this study.

**Conclusions and discussion**

While online behavior and improved self-reported measures of this behavior become increasingly important in communication research and other disciplines, our knowledge about the accuracy of these measures is still limited. For the present study a survey was conducted among the participants

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\(^3\)The same patterns were found when removing outliers, and as such the ANOVAs are not reported.
of an online panel who had a tracking software installed on their private PCs or tablets. This study extended earlier findings (e.g., Revilla et al., 2016; Scharkow, 2016) by including a wider range of explanatory factors influencing the accuracy of self-reported internet use, and including tablet
devices in addition to PCs. In addition, detailed measures of internet use and an experimental design allowed to explore how the accuracy of such measures could be improved by survey design strategies. Overall, the analysis revealed a vast gap between self-reported internet use and the equivalent measures derived from the tracking data with low correlations between both types of measures. Respondents were found to more often over-report than under-report the time they spend online. These findings correspond to previous research that has found low levels of measurement correspondence and tendencies of over-reporting for internet use (Scharkow, 2016). This pattern has also been found to apply for other media use behavior, such as the use of social media (Junco, 2013) or mobile phones (Vanden Abeele et al., 2013).

While random deviations of reported from actual behavior mainly limit the reliability of the measures and research findings, any non-random errors lead to biased results (e.g., Prior, 2009b). From a range of possible predictors factors related to actual internet use were most consistently associated to the magnitude of response errors. As has been found before, not sociodemographic or

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**Table 4. Key findings.**

<table>
<thead>
<tr>
<th>Hypothesis or Research Question</th>
<th>Accuracy</th>
<th>Over-reporting (RQ2)</th>
<th>Under-reporting (RQ2)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Factors influencing self-reporting accuracy</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Survey Interest</td>
<td>H1. The higher the level of interest in the survey, the higher the accuracy of self-reports.</td>
<td>Not supported</td>
<td>Not supported</td>
</tr>
<tr>
<td>Internet Use</td>
<td>H2. The higher the level of internet use, the lower the accuracy of self-reports.</td>
<td>Supported</td>
<td>Supported</td>
</tr>
<tr>
<td>Multitasking</td>
<td>H3. The higher the level of multitasking, the lower the accuracy of self-reports.</td>
<td>Not supported</td>
<td>Not supported</td>
</tr>
<tr>
<td>Mobile devices</td>
<td>H4. The accuracy of self-reports is lower for the use of mobile devices compared to PCs.</td>
<td>Not supported</td>
<td>Not supported</td>
</tr>
<tr>
<td><strong>Strategies to improve accuracy</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Anchoring</td>
<td>H5. Anchoring in survey questions improves the accuracy of self-reports.</td>
<td>Not supported</td>
<td>Not supported</td>
</tr>
<tr>
<td>Reference periods</td>
<td>RQ1. Which reference period (recent vs. typical) is more accurate?</td>
<td>Recent &gt; Typical Over-reporting: Recent &gt; Typical</td>
<td>No differences</td>
</tr>
<tr>
<td><strong>Discerning under- and over-reporting</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Under- and over-reporting</td>
<td>H6: Lower levels of internet use are associated to higher levels of over-reporting (a) and greater levels of internet use are associated to higher levels of under-reporting (b).</td>
<td>Supported</td>
<td>Supported</td>
</tr>
</tbody>
</table>

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**Figure 5C. Anchor–Under-reporting.**
other personality factors are most decisive for response accuracy, but factors that are closely related to
the behavior in question (Kahn et al., 2014; Revilla et al., 2016). Particularly, those with higher levels of
actual internet use as well as tablet users showed lower levels of over-reporting and, accordingly, higher
levels of under-reporting and, thus, clearly deviated from the general tendencies of over-reporting.

Related to this point and also in line with previous findings, the variance in the self-reported
measures was smaller compared to the actual behavior which could imply a lower explanatory power
of self-reported measures of internet use (Scharkow, 2016; Vanden Abeele et al., 2013; Wonneberger
& Irazoqui, 2016). Corroborating these previous studies, actual internet use was identified as a main
driver for this phenomenon: Greater levels of internet use were associated to higher levels of under-
reporting and lower levels of internet use to higher levels of over-reporting. This pattern coincides
with a more general notion about how the frequency of a behavior influences whether it is over- or
under-reported: Rare behaviors have been found to be more prone to being over-estimated while
more frequent behaviors are under-estimated more often (Schwarz & Oyserman, 2001). In other
words, vast differences of internet or other media use within the population, and more importantly
their effects, are not sufficiently reflected by self-reported data.

By discerning users of PCs and tablets, this study reveals the importance of taking the modality of
internet use or media use in general into account when assessing the accuracy of self-reports. In
contrast to previous research on more specific website use that revealed considerably lower levels of
accuracy for mobile devices (smartphones) as opposed to PC users (Revilla et al., 2016), the current
study showed that tablet users were particularly prone to under-reporting whereas PC users were
more likely to over-report. Using mobile devices “en passant” might hinder recollecting actual usage
situations. With a growing relevance of mobile devices, the problem of self-reported online media
use might thus shift from overall tendencies of over-reporting to under-reporting.

In addition to analyzing factors that may cause biases in self-reports, this study also examined two
strategies to improve the accuracy of self-reported internet use. First, different reference periods, that
is average behavior and a more recent time frame, were compared revealing higher levels of accuracy
for the more general reference period. Also for more specific forms of media use, such as news
exposure, measures of average behavior have been found to perform better compared to more recent
time frames (Althaus & Tewksbury, 2007; Chang & Krosnick, 2003).

Second, the use of anchor questions that primed the individual context of internet use (Jerit et al.,
2016; Potts & Seger, 2013) was tested with an experimental design. Triggering respondents to think
about their previous day and week was not found to facilitate more accurate estimations and reports.
A reason for this finding might be that estimating one’s duration of internet use simply is a too
complex task and improving recall is not the right starting point here. Providing population averages
as suggested, for instance, for news exposure (Prior, 2009b) might be a possible alternative. This,
however, requires a good estimate of such an average that preferably should be based on tracking
data of a sample representative for the particular group. Generally, to improve anchoring techniques
more research is necessary to, first, better understand the underlying mechanisms of providing
individual, contextual, or population cues. Second, anchor questions should be further refined and
their effectiveness assessed and compared.

**Limitations**

Probably most important for a validation study is the underlying definition of internet use that was
applied for the tracking and the survey research. In the survey—in line with common research
practice—no detailed explanation about what is considered internet use was given to the respondents
leaving ample room for subjective interpretations regarding, for instance, the type of internet use, the
minimum duration of use, or whether one has to actually look at the screen, etc. The tracking data,
in contrast, applied clear rules, such as a threshold of 120 sec as minimum duration or the
recognition of active URLs and applications. Consequently, the tracking software might recognize
internet use that is not perceived as such by a respondent, for instance, because he or she left the
room for a moment or was talking to someone else. Taking into account that respondents may overlook or ignore lengthy qualifications (Belson, 1981), providing simple and clear definitions to respondents might help to overcome subjectivity and hence improve the validity of self-reports.

Arguably, general internet use is a broad concept that comprises manifold online activities and usage motivations, such as information seeking, entertainment, or communication. However, in addition to these more specific types of usage also general use appears as a relevant explanatory concept in areas, such as uses-and-gratifications research (Papacharissi & Rubin, 2000), health communication (LaRose, Lin, & Eastin, 2003), or adolescents research (Hargittai & Hinnant, 2008). Considering that previous research indicates higher levels of accuracy for more specific online activities (Scharkow, 2016), measuring these might be the preferred option if possible in a specific research context. However, it should be noted that comparisons of the accuracy of different self-reported online activities are scarce and inconclusive. Revilla et al. (2016), for instance, found vast differences in the accuracy of self-reported visits for a range of popular websites.

The problem of a selection bias has previously been discussed for the use of tracking data and also applies to the current study (Revilla et al., 2016). To relate internet use to individual persons as opposed to households, it was necessary, for instance, to restrict the sample to single-person households. The purpose of the current study was, however, not to extrapolate findings regarding levels of accuracy or the extent of specific biases to a more general population but rather to uncover general problems of self-reported internet use and test possible solutions (see also, Mandell, 1974).

While a strength of this study was that different reference periods could be directly compared for the same respondents, the order of questions (referring first to more recent behavior, followed by typical behavior) might have also affected the response accuracy. Alternating between different reference periods might have caused confusion among the respondents and was therefore not applied.

This study started out from the assumption that tracking data can serve as a “gold standard” to assess the quality of self-reported measures of internet use. However, as all measurements also this type of data is not free of problems. For example, the tracking software itself presented limitations when it comes to the reporting of actual internet use for tablets, by which certain domains had 0 sec of activity (when they were at the end of the navigation session). While we tested all models with the actual tracking for tablets and a correction, future research should test other technologies for tracking user activity.

This was the first study comparing self-reports with tracking data of internet use to explore the accuracy of mobile device (tablet) use, in addition to internet use on home computers. That said, our data included only one (private) device per person. While all questions clearly indicated that the respondent should provide estimates considering only private use on that device (PC or tablet), this might have complicated an accurate estimation of their behavior (see Scharkow, 2016). While the design adopted by this study allowed for a clearer comparison between PC and tablet users, future research should contemplate if multiple devices per person would increase the comparability to measures typically used in survey research. Moreover, future studies relying only on tracking data should be cautious about device sharing when considering their research designs, and the tracking software to be used.

**Recommendations**

A first clear recommendations that can be derived from this study is that asking about the duration of internet use on a typical day yields more accurate reports compared to asking about yesterday and should, thus be the preferred reference period. In addition, media-use related factors appeared to be most problematic in their effects on response accuracy. First, the type of device matters with users of mobile devices showing different patterns of response errors. Second, the variance of self-reports appears smaller compared to actual internet use. It is, therefore, of crucial importance to help respondents to overcome recall and estimation problems that are related to media use, for instance, by developing anchors that are targeted at these problems.
Testing such anchors in combination with in-depth interviews would allow to better understand their effectiveness (Belson, 1981). Specifically, the effectiveness of providing population averages as anchors for duration measures should be studied. Applying multi-method approaches instead of relying on survey data only is another way to enhance the possibilities of survey and tracking data also allowing to directly assess response inaccuracies (e.g., Dvir-Gvirsman, Tsfati, & Menchen-Trevino, 2016).

In sum, while this study further contributed to our understanding of possible biases of self-reports, more strategies need to be developed and tested that help us to improve the accuracy of central measures within communication research. These include a wider range of anchors that facilitate estimation processes but also multi-method approaches, that is, combining survey data with other sources, such as tracking data, event sampling, or also qualitative data. While this study focused on measures of general internet use, research should invest in better understanding and improving more specific measures of internet use.

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**References**


