A different(ial) perspective: How social context influences the media violence-aggression relationship among early adolescents

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Chapter 3

Double dose: High family conflict enhances the effect of media violence exposure on adolescents’ aggression
ABSTRACT

We investigated how exposure to media violence and family conflict affects adolescents’ subsequent aggressive behavior. We expected a double dose effect, meaning that high media violence exposure would lead to higher levels of aggression for adolescents in high conflict families compared to low conflict families. A total of 499 adolescents (aged 10 to 14, 48% girls) participated in a two-wave longitudinal survey (4-month interval). Survey questions assessed their exposure to violence on television and in electronic games, family conflict, and aggressive behavior. Analyses revealed a significant interaction between media violence and family conflict. In families with higher conflict, higher media violence exposure was related to increased subsequent aggression. This study is the first to show a double dose effect of media violence and family conflict on adolescents’ aggression. These findings underscore the important role of the family in shaping the effects of adolescents’ media use on their social development.
Double dose: High family conflict enhances the effect of media violence exposure on adolescents’ aggression

Media violence exposure has been investigated as a risk factor for aggressive behavior for decades, with most studies demonstrating small to moderate effect sizes (Anderson et al., 2010; Bushman & Huesmann, 2006). Despite the current scientific debate about whether these effect sizes are theoretically and practically meaningful (Valkenburg & Peter, 2013b), such relatively small effects are neither strange nor unexpected. Aggression is a complex behavior caused by multiple factors (Browne & Hamilton-Giachritsis, 2005; Ribeaud & Eisner, 2010) of which media violence is only one. Research investigating the effects of media violence in concert with other predictors of aggression is needed in order to better understand its relative contribution to aggressive behavior (Bandura, 2009; Browne & Hamilton-Giachritsis, 2005).

The debate about whether and how media violence leads to aggression has received renewed attention following the tragic event at Sandy Hook elementary school in December 2012, where 20 children and six teachers lost their lives. In the aftermath of this tragedy, President Obama called for additional funding for research designed to better understand the role that media violence may play in gun violence (White House, 2013). Public policy advocates such as Common Sense Media have indicated that “one of the most pressing needs” in both the public and scientific debate is research that investigates the effects of media violence within the context of the family environment (Common Sense Media, 2013). Indeed, family conflict (defined as openly expressed anger, hostility, and aggression in the home; Moos & Moos, 1994) has been identified as a risk factor for aggressive behavior, particularly during childhood and early adolescence (Ribeaud & Eisner, 2010). Given that aggressive behavior is often conceptualized as the result of a socialization process in which both the family and the media play important roles (Bandura, 1986), investigating how media violence affects aggression within the context of family conflict is relevant (Ferguson, 2009). In this study, we respond to public and academic concerns by investigating the interactive effect of media violence and family conflict on adolescents’ aggression.

Research into media violence and family conflict
Although media violence and family conflict have both been investigated as predictors of aggression (Gentile, Coyne, & Walsh, 2011; Krahé & Möller, 2010; Tanaka, Raishevic, & Scarpa, 2010; Vandewater & Lansford, 1998), this has largely taken place in separate disciplines (Vandewater, Lee, & Shim, 2005). Media researchers have focused on the effects of media violence, while largely ignoring effects of family conflict, whereas
the reverse is true for family researchers. Although a few studies have investigated
the relative effect of media violence when controlling for family conflict (along with a
number of other risk factors; Ferguson, San Miguel, & Hartley, 2009; Huesmann, Moise-
Titus, Podolski, & Eron, 2003), it has not yet been investigated whether media violence
and family conflict have an interactive effect on aggressive behavior.

This lack of research attention for the joint effect of media violence and family conflict
on aggression is surprising. Researchers have long argued that media research should
incorporate social context as an integral part of research models rather than treat it as a
control variable (Jordan, 2004; Slater, Snyder, & Hayes, 2006). After all, media use does
not occur in a social vacuum, but rather takes place within a broader social context
that can shape and strengthen its effects (Gerbner, Gross, Morgan, & Signorielli, 1980;
Jordan, 2004; Slater et al., 2006). Furthermore, investigating interaction effects between
multiple predictors of aggression has been identified as an important goal for media
research (Valkenburg & Peter, 2013a). Doing so may show whether the relatively small
effects of media violence exposure found on an aggregate level may be larger for a
subgroup of children who are particularly vulnerable to the effects of media violence
(Valkenburg & Peter, 2013a). By ignoring relevant moderators such as family conflict,
the effect of media violence on aggression may be “diluted” across a study sample,
potentially resulting in small effects (Valkenburg & Peter, 2013a). To address this gap in
the literature, we conducted a longitudinal survey among Dutch adolescents aged 10
to 14 to test whether there is an interactive effect of media violence and family conflict
on adolescents’ subsequent aggressive behavior.

Theoretical background

Social Cognitive Theory (SCT; Bandura, 1986) has been used to explain the effects of
both media violence and family conflict on aggressive behavior (Anderson & Bushman,
Margolin & Gordis, 2000; McKelvey, Whiteside-Mansell, Bradley, Casey, Conners-
Burrow, & Barrett, 2011; Proctor, 2006). SCT states that children learn behaviors through
their own experience and through observation of others. Seeing others enact a certain
behavior and subsequently being punished or rewarded for this behavior teaches a
child which behaviors are socially acceptable and which are not. Such social cognitions
are seen as regulators of actual behavior (Huesmann & Guerra, 1997; Orue, Bushman,
Calvete, Thomaes, Oriobio de Castro, & Hutteman, 2011; Perry, Perry, & Rasmussen,
1986).

Within this social learning process, the family and the media represent two
observable role models for children and early adolescents (Bandura, 1986). Parents, in
particular, are the most important role models for youth (Anderson & Cavallaro, 2002; Bricheno & Thornton, 2007). Through their own actions, parents provide information about acceptable and unacceptable social behavior, which children can then replicate. In the case of aggression, several studies have shown that children who observe conflict within their family display more aggressive behavior themselves (Duncan, Strycker, Duncan, Okut, 2002; Farver et al., 2005; Formoso, Gonzales, & Aiken, 2000; Tanaka et al., 2010).

Given the high amount of time that adolescents spend using media (Rideout, Foehr, & Roberts, 2010), characters and behaviors observed on TV and in games also form a vital part of adolescents’ social learning process (Huesmann, 2007; Jordan, 2004). The way in which aggression is often portrayed in these media—glorified, rewarded, and performed by attractive characters (Konijn, Nije Bijvank, & Bushman, 2007)—further increases the chance that adolescents will also display such behavior. Indeed, experimental research has shown that children and adolescents who are exposed to rewarded violence or who identify with aggressive characters subsequently demonstrate increased aggressive behavior (Boyatzis, Matillo, & Nesbitt, 1995; Konijn et al., 2007).

Two theoretical models would predict that observing aggression both in the media and in the family can create a “double dose” effect (Gerbner et al., 1980). Cultivation theory includes the concept of resonance, which is the idea that people whose life experiences are congruent with what they see on television will be more affected by media messages (Gerbner et al., 1980). The Differential Susceptibility to Media Effects Model (Valkenburg & Peter, 2013a) stresses that media effects are often conditional, and that social context can reinforce the effects of exposure to media violence. Based on these theories, we expect that adolescents with high media violence exposure will become more aggressive in the context of high family conflict (double dose effect) than in the context of low family conflict. Thus, we hypothesize that family conflict strengthens the effect of media violence on aggression. Investigating this moderating role of family conflict can help us understand which adolescents are particularly susceptible to the effects of media violence exposure.

METHOD

Participants

After receiving approval from the sponsoring institution’s Institutional Review Board, a large, private survey research institute in the Netherlands (TNS NIPO) collected the data. Adolescents were recruited through TNS NIPO’s existing online panel (approximately
60,000 households) that is representative of the Netherlands. 673 Dutch adolescents between the ages of 10 and 14 years completed an online survey in January 2012. In May 2012, 499 of these adolescents agreed to participate in a second survey (i.e., 74% recontact agreement). These 499 adolescents made up the sample used in this study (85.8% sibling pairs; 47.9% girls; age at Time 1: $M = 11.87$, $SD = 1.46$).

**Measures**

**Media violence exposure**

Media violence exposure was measured using direct estimates, a measure frequently used in media research (Lee, Hornik, & Hennessy, 2008; Schmitz et al., 2004; Vandewater & Lee, 2009). Exposure to violent content on TV and in electronic games was measured with two items each (four items in total): (1) How often do you watch television programs [play games] that contain violence? and (2) On the days that you watch television programs [play games] that contain violence, how much time do you spend on this per day? Participants were given the following definition of violence: “All violence (for example, fighting and shooting) that living beings (for example, humans and monsters) do to each other.” Response categories for the first item ranged from 0 (never) to 7 (7 days per week). The second item was an open-ended question, answered by filling in hours and minutes. Items were multiplied to calculate the number of hours per week of violent television and violent game exposure. These two variables were then summed to create one variable representing violent media exposure in hours per week.

**Family conflict**

Family conflict was measured using five items from the conflict subscale of the Family Environment Scale (Jansma & Coole, 1996; Moos & Moos, 1994). Respondents were asked to indicate how often family members do the following things at home: (1) criticize each other, (2) hit each other, (3) argue, (4) curse, and (5) become so angry they start throwing things. Response categories were (1) never, (2) almost never, (3) sometimes, and (4) often. Scores were averaged to create scales ($\alpha = 0.75$), with higher scores indicating greater family conflict.

**Aggression**

Adolescents’ direct aggression was measured with six items adapted from the Direct and Indirect Aggression Scale (Björkqvist, Lagerspetz, & Kaukiainen, 1992). Adolescents were asked how often in the past six months they had done the following things to another adolescent: (1) call names, (2) push in a rough way, (3) kick or hit, (4) threaten to beat up, (5) fought with, and (6) tripped on purpose. Response categories were
(1) never, (2) 1 time in the past 6 months, (3) 2 to 3 times in the past 6 months, (4) about 1 time per month, (5) about 1 time per week, and (6) about every day. Scores were averaged to create scales ($\alpha = .83/.85$ at Time 1/Time 2), with higher scores indicating greater aggressive behavior.

**Analytic approach**

Because 86% of our sample (428 children) consisted of sibling pairs, ordinary least squares (OLS) regression was inappropriate as the assumption of independent observations is violated. This can result in over- or underestimation of coefficients due to biased estimates of standard errors (Desai & Begg, 2008; Hayes, 2006). To address this clustering in the data, multilevel modeling was used in SPSS 20. Multilevel models take into account that some variables are clustered or nested within other variables (in our case, children were nested within households). We accounted for the correlation among children within a household by allowing the mean aggression score (i.e., the intercept of the regression equation) to vary freely across households. This results in appropriate estimates of the standard errors for the regression coefficients (Desai & Begg, 2008). Parameters in this multilevel model can be interpreted in the same way as OLS regression.

All models controlled for aggressive behavior at Time 1. In addition, we investigated whether gender would be a suitable covariate in our analyses. Because of the relatively short time lag between the data collection waves, as well as the considerable skewness introduced to our data when including gender, we had some concern that the gender-included model would result in incorrect estimates. Given that gender is often included in models investigating media violence and aggression (e.g., Hopf, Huber, Weiss, 2008; Krahé, Busching, & Möller, 2012), we opted to conduct two sets of analyses: one in which gender is omitted from the model, and one in which gender is treated as a covariate. Results from both analyses are reported in the text.

The two-way interaction was probed using the Johnson-Neyman technique (Preacher, Curran, & Bauer, 2006). This technique indicates where (i.e., at which values of a continuous moderator) the effect of an independent variable on a dependent variable is significant (at $p < 0.05$). In this particular study, the Johnson–Neyman technique allows us to identify at which values of family conflict the relationship between media violence and aggressive behavior achieves statistical significance. Although the values obtained for these regions of significance are slightly less stable in multilevel models than in fixed-effects regression, this technique provides valuable information for the interpretation of interaction effects (Bauer & Curran, 2005).
RESULTS

Descriptives and intercorrelations

Table 1 presents the means, standard deviations, and correlations for all model variables. Adolescents in our sample scored relatively low on aggressive behavior, with a mean of 1.52 (SD = 0.71, scale range: 1 to 6). Family conflict was also relatively infrequent in the sample, with a reported mean of 2.10 (SD = 0.55, scale range: 1 to 4). On average, adolescents reported consuming 4.8 hours per week of media violence (SD = 8.30).

Gender correlated significantly with media violence exposure at Time 1 and aggressive behavior at Time 1 and 2. Boys were more aggressive on average than girls (boys: $M = 1.70$, $SD = 0.82$; girls: $M = 1.32$, $SD = 0.51$), and also reported about four times more media violence exposure (boys: $M = 7.60$ hours per week, $SD = 10.47$; girls: $M = 1.80$ hours per week, $SD = 2.71$).

Aggressive behavior was quite stable over time ($r = .55$, $p < 0.001$). Significant correlations were found between media violence and aggressive behavior, both cross-sectionally ($r = .27$, $p < 0.001$) and longitudinally ($r = .18$, $p < 0.001$). Family conflict was also significantly related to aggressive behavior (cross-sectional $r = .36$, $p < 0.001$, longitudinal $r = .30$, $p < 0.001$). The two predictors, family conflict and media violence, correlated at .11 ($p = 0.02$).

Table 1 Means (standard deviations) and correlations for model variables

<table>
<thead>
<tr>
<th>Variables</th>
<th>Means (SD)</th>
<th>Correlations</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>Full sample</td>
<td>Boys</td>
</tr>
<tr>
<td>1. Aggression T1</td>
<td>1.52 (0.71)</td>
<td>1.70 (0.82)</td>
</tr>
<tr>
<td>2. Media violence T1</td>
<td>4.82 (8.30)</td>
<td>7.60 (10.47)</td>
</tr>
<tr>
<td>3. Family conflict T1</td>
<td>2.10 (0.55)</td>
<td>2.13 (0.55)</td>
</tr>
<tr>
<td>4. Aggression T2</td>
<td>1.46 (0.68)</td>
<td>1.61 (0.77)</td>
</tr>
<tr>
<td>5. Gendera</td>
<td>-</td>
<td>-</td>
</tr>
</tbody>
</table>

*a* Girls = 0; boys = 1.
*p* < 0.05, ***p* < 0.001.
Main effects of media violence and family conflict

We first investigated the main effect of media violence exposure on aggressive behavior in two steps: first while only controlling for Time 1 aggression, second by adding family conflict to the model. The two predictors (Time 1 media violence exposure and family conflict) and the control variable (Time 1 aggression) were standardized in order to obtain standardized regression coefficients in the multilevel analysis. Standardized regression coefficients ($b^*$) are presented in the text; unstandardized coefficients are presented in Table 2. For the unstandardized coefficients ($b$), the predictor variables were centered at their sample mean values to reduce multicollinearity problems with their interaction terms.

Although Time 1 media violence exposure and Time 2 aggression correlated significantly (as shown in Table 1), media violence was not a significant predictor of aggressive behavior when controlling for Time 1 aggression ($b^* = 0.02, p = 0.47$). When family conflict was added to the model, the relationship between media violence and

Table 2 Unstandardized regression coefficients, standard errors, and 95% confidence intervals (CI) predicting Time 2 aggressive behavior

<table>
<thead>
<tr>
<th>Parameters</th>
<th>Main effects</th>
<th></th>
<th>Two-way interaction</th>
<th></th>
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</thead>
<tbody>
<tr>
<td></td>
<td>$b$</td>
<td>$SE$</td>
<td>95% CI</td>
<td>$b$</td>
</tr>
<tr>
<td><strong>Regression coefficients (fixed effects)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Intercept</td>
<td>1.46***</td>
<td>0.03</td>
<td>[1.40; 1.51]</td>
<td>1.45***</td>
</tr>
<tr>
<td>Aggression Time 1</td>
<td>0.47***</td>
<td>0.04</td>
<td>[0.39; 0.55]</td>
<td>0.46***</td>
</tr>
<tr>
<td>MVE</td>
<td>0.00</td>
<td>0.00</td>
<td>[−0.00; 0.01]</td>
<td>0.00</td>
</tr>
<tr>
<td>FC</td>
<td>0.15**</td>
<td>0.05</td>
<td>[0.05; 0.25]</td>
<td>0.14**</td>
</tr>
<tr>
<td>MVE * FC</td>
<td></td>
<td></td>
<td></td>
<td>0.01*</td>
</tr>
<tr>
<td><strong>Variance components (random effects)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Random intercept</td>
<td>0.04†</td>
<td>0.02</td>
<td>[0.01; 0.12]</td>
<td>0.04†</td>
</tr>
</tbody>
</table>

Note: MVE = Media Violence Exposure; FC = Family Conflict. For ease of readership, Table 2 reflects the analyses without gender.

† $p < 0.10$, ‡ $p < 0.05$, § $p < 0.01$, *** $p < 0.001$. 
aggression was unchanged ($b^* = 0.02, p = 0.48$). Family conflict did predict subsequent aggression ($b^* = 0.08, p = 0.003$), with increased family conflict predicting increased aggression. These results did not change when we added gender to the model as covariate: Media violence was not associated with subsequent aggression ($b^* = −0.00, p = 0.89$) whereas family conflict was ($b^* = 0.08, p = 0.002$).

**Interaction between media violence and family conflict**

Following main effects testing, we tested our hypothesis that media violence exposure leads to more aggression in the context of high family conflict (i.e., double dose) than in the context of low family conflict. Results supported the double dose effect. There was a significant two-way interaction between media violence and family conflict ($b^* = 0.06, p = 0.01$). The Johnson–Neyman technique indicated that the effect of media violence on aggression was significant only at higher values of family conflict, that is, for scores of 2.44 or higher on the four-point family conflict scale. As shown in Figure 1, high media violence exposure in combination with higher levels of family conflict led to increased aggression—providing evidence for a double dose effect.

The double dose effect was also found when we included gender as covariate in the model. The interaction between media violence and family conflict remained significant ($b^* = 0.06, p = 0.02$). The effect of media violence on aggression was again found for higher levels of family conflict (in this analysis, for scores of 2.95 and higher on the family conflict scale). The Johnson–Neyman technique also indicated a significant effect of media violence on aggression for low family conflict (i.e., scores of 1.37 and lower on a scale ranging from 1 to 4). For low conflict families, the pattern observed was the reverse: Increased media violence was related to decreased aggression. However, we are cautious to interpret this result, as it may reflect an artifact of the data given (a) the conservative nature of the gender-controlled analyses, (b) the limited range of this effect (0.37 data points on the 4-point scale), and (c) the low number of adolescents with high media violence and low family conflict scores ($n = 4$).
DISCUSSION

This study investigated the interactive effect of media violence exposure and family conflict on adolescents’ aggressive behavior. We found support for the hypothesized double dose effect: Adolescents’ aggression increased when they were exposed to both high media violence and high family conflict. This finding provides three important implications for research and practice. First, the presence of a double dose effect is consistent with the idea that not all media consumers are affected by media violence in the same way (Valkenburg & Peter, 2013a). Although many studies have shown main effects of media violence on aggression (e.g., Gentile et al., 2011; Huesmann et al., 2003; Krahé & Möller, 2010), our study shows that media violence exposure may have a stronger effect on adolescents from particular social contexts. By identifying and testing when media violence may have a stronger effect on aggression, rather than expecting equal effects for all media consumers, our study contributes to a more nuanced understanding of the effects of media violence. Further research that seeks to investigate other potential moderators of the media violence-aggression relationship would be worthwhile.

Figure 1  High media violence exposure and high family conflict lead to increased aggression in adolescents (* = significant at p < 0.05). Note: low, mean, and high family conflict represent values of 1.76, 2.10, and 2.44 on the family conflict scale.
Second, our findings highlight the importance of incorporating social context in media research. Adolescents living in higher conflict families showed increased aggressive behavior as a result of higher media violence exposure. One explanation for such a double dose effect is that more frequent experiences with aggression—both directly via family members and indirectly via the media—lead to more accessible aggression-related scripts in adolescents’ memory (Farver et al., 2005; Huesmann, 2007). Having more readily accessible aggressive scripts increases the likelihood of using aggressive behavior in social interactions (Lösel, Bliesener, & Bender, 2007). A second way in which frequent exposure to aggression in both the family and the media may lead to increased aggression is via adolescents’ social norms. High media violence consumers in high conflict families can learn from two social contexts that aggression is acceptable social behavior (Orue et al., 2011). Such normative beliefs have been shown to predict increased aggression (Henry, Guerra, Huesmann, Tolan, VanAcker, & Eron, 2000; Huesmann & Guerra, 1997; Perry et al., 1986). Finally, a double dose effect may be explained by maladaptive processing of social information. Research has shown that children living in family environments characterized by high levels of negative emotional expression are more likely to experience maladaptive processing styles and subsequent conduct problems (Schultz & Shaw, 2003). Maladaptive processing styles are operationalized as hostile attribution bias (i.e., a tendency to view others’ intentions as mean) and maladaptive response generation (i.e., when presented with a negative situation, children generate aggressive response solutions as opposed to more prosocial response options). It may be that children growing up in households with high family conflict and high media violence develop maladaptive processing styles and, as a result, demonstrate increased aggression. Currently, it is unknown which of these three mechanisms—aggressive scripts, normative beliefs, or maladaptive processing—may explain this double dose effect. Research which seeks to identify how a combination of real-life aggression and media aggression leads to increased aggressive behavior in adolescents would advance our understanding of this effect.

Third, our findings provide a starting point for practitioners who work on the prevention or reduction of adolescents’ aggressive behavior. Given that family conflict not only predicted aggression individually, but also strengthened the effect of media violence on aggression, it seems reasonable that most ground can be gained by reducing aggression within the family. Family members’ aggressive behavior may have a more pronounced influence on adolescents’ aggression compared to media violence due to their proximity and emotional closeness to the adolescent (Ferguson, 2009). Helping family members in high conflict families to recognize and change their
norms about acceptable and unacceptable social behavior would be a fruitful first step towards reducing adolescents’ aggressive behavior. However, it is important to recognize that changing household norms and reducing family conflict may not always be easy to achieve, particularly in high conflict families. For these families, our results speak to the benefit of reducing media violence as an alternative way of decreasing adolescents’ aggressive behavior.

Limitations
It is important to recognize that the effect sizes found in our study were relatively small. In fact, despite a significant bivariate correlation between media violence and aggression, this study did not find a main effect of media violence exposure on aggression—a finding inconsistent with previous research (e.g., Gentile et al., 2011; Huesmann et al., 2003; Krahé & Möller, 2010). Although small effects are not uncommon in media research (Valkenburg & Peter, 2013b), three methodological choices should be taken into account when interpreting our results. First, our choice of media violence exposure measure may have resulted in underestimation of the effect sizes found. This measure asked adolescents to report how often and how long they consume violent TV shows and games during the week. Although such direct estimates have been validated for use among adolescents (Schmitz et al., 2004; Van der Voort & Vooijs, 1990), the estimation procedures involved in these measures can be difficult for adults and adolescents alike, potentially resulting in over- or underestimation of their actual exposure (Robinson & Godbey, 1997). Consequently, the small but significant effect sizes found in our study, as well as the absence of a significant main effect of media violence on aggressive behavior, may be partly due to our choice of media violence measure.

Second, the time lag between waves in our study was relatively short (four months). Since aggressive behavior was quite stable in our study (see the longitudinal correlation between Time 1 and Time 2 aggression in Table 1), this time lag may have been too short to show larger effects of media violence and family conflict on aggression. Third, our sample consisted of typically-developing adolescents who scored relatively low on media violence, family conflict, and aggressive behavior. Because of the restricted range of scores on these variables, the relationships between them may have been attenuated. Still, despite these caveats, we found a double dose effect of media violence and family conflict on adolescents’ subsequent aggression, which speaks to the strength of this effect. Future studies may find this effect to be stronger when investigating it in a study with a longer time lag and an at-risk sample in which media violence, family conflict, and aggressive behavior are more frequent. Furthermore,
replication of our study using a measure of media violence exposure that relies less on adolescents’ estimates of media use across longer periods (such as the Favorites measure by Anderson and Dill, 2000) would be an appropriate next step.

Conclusion
This study is the first to investigate the interactive effect of media violence and family conflict on adolescents’ aggression. Results showed that media violence may be a stronger risk factor for adolescents in high conflict families compared to those in low conflict families. Our findings are a first step towards improving the public and scientific understanding of whether and how media violence leads to aggression in the context of other risk factors (Common Sense Media, 2013; Tanaka et al., 2010; Valkenburg & Peter, 2013a). Practically speaking, these findings underscore the important role of the family in the development of early adolescents’ social behavior. It is important for parents to realize that their home environment can strengthen the effect of media violence on their child’s aggressive behavior. Fostering discussions about what constitutes acceptable social behavior, both in the media and in real life, can be a useful way for parents to mitigate potential negative effects of violent media exposure.