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Validation of the Amsterdam Beverage Picture Set: A Controlled Picture Set for Cognitive Bias Measurement and Modification Paradigms

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Background: Alcohol research may benefit from controlled and validated picture sets. We have constructed the Amsterdam Beverage Picture Set (ABPS), which was designed for alcohol research in general and cognitive bias measurement and modification in particular. Here, we first formulate a position on alcohol stimulus validity that prescribes that alcohol-containing pictures, compared to nonalcohol-containing pictures, should induce a stronger urge to drink in heavy drinkers than in light drinkers. Because a perceptually simple picture might induce stronger cognitive biases but the presence of a drinking context might induce a stronger urge to drink, the ABPS contains pictures with and without drinking context. By limiting drinking contexts to simple consumption scenes instead of real-life scenes, complexity was minimized. A validation study was conducted to establish validity, to examine ABPS drinking contexts, and to explore the role of familiarity, valence, arousal, and control.

Methods: Two hundred ninety-one psychology students completed the Alcohol Use Disorders Identification Test, as well as rating and recognition tasks for a subset of the ABPS pictures.

Results: The ABPS was well-recognized, familiar, and heavy drinkers reported a greater urge to drink in response to the alcohol-containing pictures only. Alcohol presented in drinking context did not elicit a stronger urge to drink but was recognized more slowly than alcohol presented without context.

Conclusions: The ABPS was found to be valid, although pictures without context might be preferable for measuring cognitive biases than pictures with context. We discuss how an explicit approach to picture construction may aid in creating variations of the ABPS. Finally, we describe how ABPS adoption across studies may allow more reproducible and comparable results across paradigms, while allowing researchers to apply picture selection criteria that correspond to a wide range of theoretical positions. The latter is exemplified by ABPS derivatives and adoptions that are currently under way.

Key Words: Implicit Cognitions, Cognitive Bias Modification, Alcohol, Pictures, Standardization.

Dual Process Models of addiction attempt to explain why people persevere in addictive behaviors while consciously knowing that the long-term consequences may be dire, by proposing that this perseverance is partly driven by implicit or automatically activated cognitive processes that are biased “toward” drug consumption (Stacy and Wiers, 2010). Compared to explicit processes, implicit processes can be characterized by being relatively inaccessible to conscious reflection or control and being relatively fast (De Houwer et al., 2009). Implicit cognitive processes can be measured with several indirect or implicit tasks. For example, an approach bias can be assessed with the Stimulus Response Compatibility task (De Houwer et al., 2001) and with the Approach Avoidance Task (Rinck and Becker, 2007), implicit memory associations with the Implicit Association Task (Greenwald et al., 1998), and attentional bias measured with several indirect or implicit tasks. For example, an approach bias can be assessed with the Visual Probe Test (VPT; MacLeod et al., 1986). In addition, recent research has demonstrated that these processes may be directly targeted with interventions that are collectively called Cognitive Bias Modification (CBM). In some studies, CBM has shown promise to improve effectiveness of interventions to abstain or reduce use (for a review, see Wiers et al., 2013).

In these measurements and interventions, pictures of beverages containing alcohol (henceforth called “alcohol pictures”) and pictures of beverages that do not contain alcohol (henceforth called “nonalcohol pictures”) are presented to participants in various computerized procedures. A multitude of tasks exist to measure implicit processes in addiction.
We state that relative to alcohol, a valid nonalcohol picture should elicit a relatively weaker, but not a negative urge to drink in heavy drinkers compared to light drinkers. Via the latter criterion we emphasize that the variation in urge to drink caused by alcohol and nonalcohol pictures is due to an appetitive response toward alcohol. If, for example, heavy drinkers would show a weaker urge to drink in response to nonalcohol than light drinkers would, but a similar urge in response to alcohol, then variation in the pictures would indeed cause variation in urge to drink, but this would more likely be due to aversive responses toward the nonalcohol than due to appetitive responses toward alcohol. Finally, note that these hypotheses assume that the attribute “urge to drink” is more easily influenced in heavy drinkers than light drinkers.

Besides urge to drink, pictures were rated on various aspects that have been used in prior validity research. Behaviors were rated on familiarity (Pulido et al., 2010; Tibboel et al., 2010) and both self-reported and forced-choice recognition (Grüsser et al., 2000; Pulido et al., 2010). Additionally, pictures were rated on each of the 3 dimensions outlined in the International Affective Pictures System (Lang et al., 2005): valence, arousal, and control. Familiarity, recognition, and affective value are not used to assess validity, but we do acknowledge their relevance by providing summary scores on all these measures for each of the ABPS pictures. This way, a researcher using the ABPS can apply his preferred picture selection criteria, for instance by equalizing or maximizing differences in affective values.

The ABPS has been designed to be particularly suitable for CBM research. A picture feature that appears to be particularly relevant to CBM is the presence or absence of a drinking context. As yet, the absence or presence of a drinking context has mainly been operationalized as bland backgrounds or real-life scenes (wherein real-life is an umbrella term for any nonstudio picture, such as a picture shot inside of a bar or at a party), respectively. In heavy drinkers, self-reported craving (Lee et al., 2006) was found to be more strongly affected by a beverage in a drinking context than a beverage in front of a bland background, giving rise to the suggestion that real-life scenes seem most effective in eliciting an urge to drink. Note that Nees and colleagues (2012) have drawn a similar conclusion based physiological measures of valence and arousal instead of self-reported craving. However, real-life scenes are also more complex; findings with eye fixation times (but less so with RT) have led to the suggestion that complex pictures may require more cognitive processing, and so interfere with visual probe like attentional bias measures (Miller and Fillmore, 2010). Hence, while a real-life scene may be most effective for triggering urges to drink (Lee et al., 2006), a bland background may be most effective for assessing attentional bias (Miller and Fillmore, 2010).

The ABPS features beverages in front of a bland background with or without drinking attributes or a person. Each of the beverages was displayed in 6 contexts of which 3 were passive (only displaying the beverage) and 3 were active (displaying the beverage being served, opened, or held/con-
The active pictures were aimed at providing a drinking context that may trigger a relatively strong urge to drink, while keeping complexity to a minimum, thus facilitating the assessment of implicit cognitive processes such as attentional bias. To test the hypothesis that active pictures and passive pictures are similarly perceptually complex, in the current study, recognition speeds were compared across drinker groups and contexts.

In summary, we expected that the alcohol pictures induced a stronger urge to drink in heavy drinkers than in light drinkers. We expected that the nonalcohol pictures induced a relatively weaker, but not negative urge to drink in heavy drinkers compared to light drinkers. Examining the effects of context, we expected a stronger urge to drink in heavy drinkers in response to active alcohol pictures than passive alcohol pictures, but similar recognition speeds. Exploratively, we examined the role of familiarity, recognizability as alcohol or nonalcohol, and affective values in terms of valence, arousal, and control.

MATERIALS AND METHODS

Participants

Participants were 291 first-year psychology students from the University of Amsterdam that took part in the study for course credits. Only participants that had drunk any alcohol the past year were included.

Design

Each participant completed 96 trials in a recognition and rating task. During each trial 1 picture of the ABPS was presented. Each participant viewed 8 alcohol and 8 nonalcohol beverages, displayed in each of the 3 active contexts (with beverage in a full bottle, full bottle with empty glass, and bottle with full glass) and each of the 3 passive contexts (with beverage being served, opened, and held/consumed). For the 8 alcohol beverages, 2 exemplars were picked randomly from each of the 4 alcohol types (beer, wine, spirits, and alcopops). For the nonalcohol beverages, exemplars were picked randomly from each of the 3 non-alcohol beverage types (soda, water, and miscellaneous) such that 2 types were represented by 3 beverages and 1 type with 2, making a total of 8.

Measures

Alcohol Use Disorders Identification Test. The Alcohol Use Disorders Identification Test (AUDIT; Babor et al., 2001; Saunders et al., 1993) was used to distinguish heavy from light drinking. In the general population a cutoff of 8 has been recommended to identify problematic drinking (Babor et al., 2001). In students, cutoffs as high as 11 have been recommended (Fleming et al., 1991). Given the prevalence of heavy drinking in the current sample (mean AUDIT was 10.1), relatively heavy and light drinkers were distinguished via median split.

Recognition Task. In the recognition task, participants pressed the “Q” key on the keyboard if a picture showed a beverage containing alcohol, and the “P” key if a picture showed a beverage without alcohol. The mapping of keys to alcohol and nonalcohol responses
was counterbalanced across participants. Participants were requested to respond as quickly and accurately as possible. Each trial started with a gray fixation cross presented for 800 ms, after which it turned black for 200 ms. Next, the beverage picture was presented for 500 ms. After each response, the screen was blank for 500 ms, until the next trial appeared. For each participant, proportion of correct responses and mean RTs of correct responses were calculated from task results.

**Rating Task.** In the rating task, a picture was displayed at full size for 1.5 seconds, after which the picture zoomed out while on the side a panel appeared with 6 items. Three items concerned the affective ratings for valence, arousal, and control. For each of these items, 9 answer options were provided, with Self-Assessment Manikins above every other answer option (valence manikins based on Suk, 2006, p. 206; arousal and control based on Lang, 1980). Answers were scored from −4 on the left side of the scale to 4 on the right side, wherein negative values represented more negative affect, relaxation, and being controlled, while positive values represented more positive affect, excitement, and having control. Urge to drink was assessed via the question: “How much would you like to drink this drink today?” with a 9-point bipolar scale ranging from “absolutely not” (scored −4) to “very much” (scored 4). Finally, 2 items assessed familiarity by inquiring whether the participant had ever consumed this beverage and whether the participant knew this beverage. The latter 2 items were answered dichotomously; yes or no. Mean ratings of desire to drink and affective responses were calculated, as well as proportion of yes answers on familiarity items.

**Procedure**

The validation study was part of a larger battery of psychological tests, administered in 4 sessions, with 1 ensuing each week. The sample was randomly split in 3 groups, with participants in each group simultaneously taking part in each session in a large classroom. During the first session the AUDIT was administered. During each session, 24 pictures were rated, randomized such that each session contained an equal number of alcohol and nonalcohol pictures. Participants first performed the recognition task and then the rating task.

**Data Analysis**

First, participants were divided into a light drinking and heavy drinking group based on their AUDIT scores. Response times in recognition task, and desire to drink and affective ratings in rating task were subjected to mixed analyses of variance (ANOVAs) with between-subject factors: drinking (light drinkers vs. heavy drinkers), and within-subject factors: alcohol (alcohol vs. nonalcohol beverage) and context (active vs. passive). Hypotheses were formulated in terms of main and interaction ANOVA effects, with corresponding follow-up tests if interactions were significant with an alpha of 0.05. To compare our categorization of heavy and light drinkers with an interpretation of AUDIT score as continuous variable, main hypotheses on urge to drink and explorations on affective value were also examined via Pearson correlations and linear regressions. Percentages correct in recognition task and percentages of yes on familiarity items were generally high, so these were subjected to Wilcoxon and Mann–Whitney tests.

**RESULTS**

**Descriptives**

Of the 291 participants that completed the experiment, 12 were dropped for reporting not having consumed any alcohol the past year in the AUDIT. Of the remaining participants, 193 were female, while 86 were male. Age ranged from 17 to 35 years (M = 19.5, SD = 2.2). Table 1 shows descriptives of each of the measures administered during the recognition and rating tasks. AUDIT scores significantly differed between men and women, t(277) = 4.56, p < 0.001, d = 0.59; men had a mean AUDIT score of 12.0 (SD = 4.7), and women a score of 9.2 (SD = 4.8). Based on items 1 and 2 of the AUDIT we constructed a measure for average number glasses of alcohol consumed every week. Men (M = 14.0, SD = 11.0) and women (M = 7.6, SD = 6.9) differed significantly in weekly drinking. Weekly alcohol consumption was strongly correlated with AUDIT score (r = 0.69, p < 0.001). Participants were classified as heavy or light drinkers based on the overall median of AUDIT scores, which was 10. Because of structural differences of AUDIT scores between men and women, the split yielded unequal numbers of men and women in the heavy and light drinker groups: For men, 63 were classified as heavy and 23 as light, while for women, 81 were classified as heavy and 112 as light. To test the validity of the median split across genders, all succeeding ANOVAs included gender as factor, the effects of which have been reported in a separate section.

**Picture Validity**

To test whether heavy drinkers felt a stronger urge to drink in response to alcohol pictures than light drinkers, we examined the interaction between drinker group (heavy vs. light) and beverage (alcohol vs. nonalcohol) on urge to drink. This interaction was significant, F(1, 275) = 15.9, p < 0.001, η²p = 0.055. In line with our hypothesis, heavy drinkers reported a stronger urge to drink (M = −0.42, SD = 1.42) in response to the alcohol pictures than light drinkers did (M = −1.39, SD = 1.46), F(1, 275) = 14.8, p < 0.001, η²p = 0.051. Simple comparisons between heavy and light drinkers revealed this difference to have a Cohen’s d of 0.677. Also in line with our hypothesis, heavy drinkers did not report a stronger urge to drink in response to the nonalcohol

### Table 1. Descriptives of the Measures Administered in the Rating and Recognition Tasks

<table>
<thead>
<tr>
<th>Task measure</th>
<th>Recognition</th>
<th>Rating</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Recognition</td>
<td>Rating</td>
</tr>
<tr>
<td>RT in ms</td>
<td>Proportion correct</td>
<td>Urge to drink</td>
</tr>
<tr>
<td>mean (SD) 639.26 (87.27)</td>
<td>0.94 (0.04)</td>
<td>0.10 (1.16)</td>
</tr>
</tbody>
</table>

RT, response time.
responded faster to alcohol than nonalcohol pictures, contexts revealed that in the active context, participants = drinkers in response to alcohol and nonalcohol pictures during rating task. ages were overall high. In the recognition task, heavy drinkers attained a slightly lower percentage correct (M = 94.5%) for alcohol pictures in the recognition task than light drinkers did (M = 95.6%), U = 8,171, p = 0.019, Z = 2.34, but not so for nonalcohol pictures, U = 9,677, p = 0.949. To examine familiarity, we inspected percentage of beverages known and consumed in the rating task. Heavy drinkers (M = 88.3%) reported having consumed a larger number of the alcohol beverages in the ABPS than light drinkers (M = 80.0%), U = 6,404.5, p < 0.001, Z = 4.95, which was not so for nonalcohol beverages, U = 9,625.5, p = 0.874. Heavy drinkers reported knowing a similar number of the alcohol beverages as light drinkers did, U = 9,102.5, p = 0.328, which was also the case for nonalcohol beverages, U = 9,173.5, p = 0.272.

Familiarity and Recognizability

To examine recognition, we inspected percentages correct in the recognition task. As shown in Table 1, these percentages were overall high. In the recognition task, heavy drinkers attained a slightly lower percentage correct (M = 94.5%) for alcohol pictures in the recognition task than light drinkers did (M = 95.6%), U = 8,171, p = 0.019, Z = 2.34, but not so for nonalcohol pictures, U = 9,677, p = 0.949. To examine familiarity, we inspected percentage of beverages known and consumed in the rating task. Heavy drinkers (M = 88.3%) reported having consumed a larger number of the alcohol beverages in the ABPS than light drinkers (M = 80.0%), U = 6,404.5, p < 0.001, Z = 4.95, which was not so for nonalcohol beverages, U = 9,625.5, p = 0.874. Heavy drinkers reported knowing a similar number of the alcohol beverages as light drinkers did, U = 9,102.5, p = 0.328, which was also the case for nonalcohol beverages, U = 9,173.5, p = 0.272.

Affective Value

To examine affective values of the pictures, we tested interactions between drinker group (heavy vs. light) and beverage (alcohol vs. nonalcohol) on affective ratings. No such interactions were found, for valence, F(1, 275) = 2.16, p = 0.143, nor arousal, F(1, 275) = 0.371, p = 0.543, nor control, F(1, 275) = 0.060, p = 0.806. In contrast with ANOVA results, continuous AUDIT scores were significantly correlated with the average difference in valence rating between alcohol and nonalcohol pictures, r = 0.22, p < 0.001. This effect remained in a linear regression of difference in valence rating on AUDIT score, gender, and AUDIT × gender interaction, in which the coefficient for AUDIT score was significant, B = 0.13, t = 5.23, p < 0.001, n^2 = 0.09. In line with ANOVA results, AUDIT scores were not significantly correlated with arousal, r = 0.10, p = 0.09, or control, r = −0.06, p = 0.36.

Gender Differences

To examine gender differences, we tested interactions between drinker group (heavy vs. light), beverage (alcohol vs.
nonalcohol), and gender (male vs. female) on all continuous dependent variables (i.e., urge to drink, RTs, and affective responses). No such interactions were found, indicating that responses of heavy and light drinkers to alcohol and nonalcohol pictures were similar across men and women. Note that there were various effects of gender regardless of drinker group. For urge to drink, there was an interaction between gender and beverage, $F(1, 275) = 9.85, p = 0.002, n^2_p = 0.035$. Follow-up analyses revealed that, controlling for drinker group, men ($M = -0.35, SD = 1.58$) reported a stronger urge to drink than women ($M = -1.13, SD = 1.43$) in response to alcohol pictures, $F(1, 275) = 7.62, p = 0.006$, but men and women reported similar desire in response to nonalcohol pictures, $F(1, 275) = 0.887, p = 0.347$. For valence, there was also an interaction between gender and beverage, $F(1, 275) = 7.40, p = 0.007, n^2_p = 0.026$. Follow-up analyses revealed that, controlling for drinker group, women ($M = 2.11, SD = 1.25$) reported more positive valence than men ($M = 1.61, SD = 1.39$) in response to nonalcohol pictures, $F(1, 275) = 5.81, p = 0.017, n^2_p = 0.021$, but men and women reported similar valence in response to alcohol pictures, $F(1, 275) = 0.065, p = 0.799$. There were no significant effects of gender on arousal, control, or RT.

**DISCUSSION**

We have introduced the ABPS, a stimulus set aimed to be useful for alcohol research in general, and for cognitive bias measurement and modification in particular. The main aim of the current study was to examine the validity of the ABPS as well as the perceptual simplicity of its pictures with and without a drinking context. Additionally, we explored recognizability, familiarity, affective value, and given the differences in drinking behavior between men and women in our sample of college students, differences between genders. Below, we reflect on each of these aims, review our approach for distinguishing between heavy and light drinker groups, limitations of the student sample, and our conceptualization of validity. Finally, we provide a short overview of future developments of the ABPS.

In line with our validation criteria, heavy drinkers reported a stronger urge to drink in response to ABPS pictures that contained alcohol than light drinkers did. Also, as expected, pictures featuring beverages without alcohol elicited a weaker, but not negative difference in urge between heavy and light drinkers. Confirming these findings, in a continuous interpretation of AUDIT score, higher scores were related to higher reported urge to drink in response to alcohol relative to nonalcohol. Thus, given the assumption that “urge to drink” is more easily influenced in heavy drinkers than light drinkers, we conclude that: (i) the ABPS provides valid alcohol stimuli, since variation in the pictures (alcohol vs. nonalcohol) cause variation in urge to drink and (ii) this is due to an appetitive response to the alcohol pictures.

The ABPS was designed to be suitable for cognitive bias measurement and modification. Accordingly, the pictures were developed to be simple, with a minimum of distracting elements. To provide a drinking context, half of the pictures displayed a beverage being operated on (active pictures), while half did not (passive pictures). In line with Lee and colleagues (2006), it was hypothesized that active pictures, by displaying a drinking context, would induce a stronger desire in heavier drinkers. In contrast with Miller and Fillmore (2010), it was hypothesized that active pictures were similarly complex as passive pictures, so active pictures would not be recognized more slowly than passive pictures. Both hypotheses were disconfirmed: active pictures did not induce a stronger urge to drink in heavy drinkers than in light drinkers, and active pictures were recognized more slowly than passive pictures. Hence, the ABPS active pictures did not succeed in being more desirable, nor in being similarly complex as the passive pictures. Based on these findings, the conclusions of Lee and colleagues (2006) could be further qualified, as it might be the social nature of their contexts that is essential to inducing a stronger urge to drink in heavy drinkers. Given the findings on recognition speed, we speculate that not only real-life scenes (Miller and Fillmore, 2010), but any context at all, may introduce complexity to such a degree that it may attenuate the assessment of attentional bias. However, it should be noted that there were no systematic differences in how heavy and light drinkers responded to the presence and absence of drinking context. Additionally, no measures of implicit cognition were administered, so any conclusions drawn are tentative at best. The current pattern of results was that beverage pictures without context were recognized fastest and most consistently across alcohol and nonalcohol categories, so tentatively, these may be most suitable for implicit cognition measurements.

Exploring recognizability and familiarity, we found that ABPS beverages were correctly recognized as containing alcohol and nonalcohol and were familiar. ABPS alcohol pictures were rated as more familiar but recognized worse by heavy drinkers than light drinkers, while nonalcohol pictures were similarly familiar and well-recognized across heavy and light drinkers. Perceived familiarity may be a confound in sensitive paradigms such as attentional blink (Tibboel et al., 2010) or when figure-ground asymmetry needs to be considered (Rothermund and Wentura, 2001). However, differences found in this study were relatively small. The online Appendix allows controlling for possible confounds by preselecting stimuli for universal familiarity across drinker groups.

Exploring affective values of the pictures, it was found that across heavy and light drinkers, affective values, both for valence, arousal, and control, were relatively similar for alcohol and nonalcohol pictures. However, when interpreting AUDIT score continuously, higher AUDIT scores were found to be related to higher valence ratings in response to alcohol relative to nonalcohol. We carefully conclude that the ABPS alcohol and nonalcohol pictures have a relatively constant affective value. From a cue reactivity perspective, such consistency may be a desirable characteristic (Grüßer
et al., 2002; Pulido et al., 2010). From an incentive sensitization perspective, it is interesting to note that results were not consistent with wanting as indicated by higher arousal and only weakly consistent liking as indicated by higher valence (Grüsser et al., 2002; Nees et al., 2012). One explanation for this difference between the findings with the ABPS and other picture sets might be the absence of real-life scenes with social components in the ABPS. As indicated by Nees and colleagues (2012), real-life scenes may contain unique qualities that induce a differential affective response in heavy drinkers. Another explanation is based on the way heavy and light drinkers were distinguished.

Participants generally scored rather high on the AUDIT, as the majority would have been classified as hazardous drinkers when using the commonly used cutoff of 8 (Cate-tano and Babor, 2006). As an alternative cutoff, a median split was used with a cutoff of 10. Such high cutoffs have indeed been used in student populations (Fleming et al., 1991; Johnsson and Berglund, 2006; Murphy and Garavan, 2011). Since there were structural differences in AUDIT scores between men and women, AUDIT groups were confounded with gender, such that the heavy drinker group contained a higher proportion of men than the light-drinking group did. However, we do believe the distinction between heavy and light drinkers made in this study via median splits distinguished between 2 degrees of severity of alcohol use in a manner that was consistent across genders. Findings that confirm this belief were that men and women showed systematic differences both on urge to drink and valence in response to alcohol and nonalcohol pictures, but none of these differences were moderated by drinker group. Also, findings with continuous interpretations of AUDIT score were relatively consistent with dichotomized AUDIT scores. Nevertheless, while the large sample of the current study provided relatively high statistical power, the use of a median split may have reduced it somewhat. Indeed, prior studies in college samples did find a positive association between drinking and valence ratings of alcohol pictures when using quantitative drinking data (Pulido et al., 2009), but not with a dichotomized AUDIT score (Billieux et al., 2011).

We have formulated a position on stimulus validity based on the theory of Borsboom and colleagues (2004), by proposing that stimuli are valid for influencing an attribute if variation in the stimuli cause variation in the attribute. We chose “urge to drink” as our criterion attribute because we presupposed this criterion to be relatively uncontroversial and universal. However, there are some conceptual and methodological issues with our position and study design that as yet are unresolved. One issue is that this criterion is not specific for alcohol. More precisely, by comparing urge to drink in response to alcohol and nonalcohol pictures we compare “urge to drink alcohol in response to alcohol pictures” to “urge to drink nonalcohol in response to nonalcohol pictures.” Nonalcohol pictures might well induce an urge to drink alcohol as well, for instance when a heavy drinker consistently drinks alcohol mixed with soda. Hence, a more accurate attribute and corresponding measure might be “urge to drink alcohol.” However, we do believe this does not invalidate our finding; if the nonalcohol pictures would indeed induce a stronger urge to drink alcohol in heavy drinkers, then heavy drinkers would also have reported a stronger urge to drink in response to the nonalcohol pictures. The latter was not the case.

Another limitation concerns our choice of measures in light of the primary aim of the ABPS. The ABPS was primarily developed for research of implicit cognitions, while we used a direct measure to validate the picture set. One could presuppose that a picture of which heavy drinkers do not report that it triggers a stronger urge to drink might nevertheless trigger an attentional bias. However, such a presupposition risks conflating the validity of a measure with the validity of a stimulus. To exemplify, if a VPT using a particular stimulus set would not find a greater attentional bias in heavy drinkers, it is unclear whether this invalidates the existence of alcohol attentional bias, the validity of the VPT as a measure of attentional bias, or the validity of the stimulus set in causing attentional bias. Hence, while validation via a direct measure might be suboptimal, it does prevent the risk of such a conflation.

A more comprehensive assessment of stimulus validity would require comparisons across paradigms. The ABPS (Van Deursen et al., 2013) and its Italian version (Boffo et al., 2015) are currently being used in a cognitive bias measurement and modification study, allowing current findings to be compared with measures of implicit cognition. Furthermore, these studies may answer whether noncontextualized pictures are indeed most effective to measure and train cognitive biases. Additionally, variations of the ABPS have been constructed for food, tobacco, and cannabis (ongoing studies), using a similar construction procedure as outlined in the online Appendix, where both instructions for construction of such a stimulus set and the stimulus set itself can be found.

In conclusion, the ABPS appears to be a valid picture set for CBM research and application in young adults. The controlled nature of the ABPS is useful in cognitive bias measurement and modification paradigms, especially for feature matching as practiced in visual probe tasks (Bradley et al., 2003), while the large number and variety of beverages allows for stimulus personalization (Fadardi and Cox, 2009; Houben and Wiers, 2007; Tapert et al., 2003) and multisession training (Eberl et al., 2013; Van Deursen et al., 2013; Wiers et al., 2011). Results on stimulus context might indicate that the real-life characteristics of social scenes might be essential to both induce a stronger urge to drink as well as wanting and liking. Hence, for optimal assessment of cue-reactivity in relation to physiological reactions or brain responses, pictures with real-life scenes such as those featured in the Geneva Appetitive Alcohol Pictures (Billieux et al., 2011) and the stimulus set of Pulido and colleagues (2010) may be more suitable. However, any context may delay
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References


**SUPPORTING INFORMATION**

Additional Supporting Information may be found in the online version of this article:

- Appendix S1. ABPS manual.
- Appendix S2. ABPS descriptions.
- Appendix S3. ABPS pictures.